Moderators of interventions designed to enhance physical and psychological functioning among younger women with early-stage breast cancer

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Moderators of Interventions Designed to Enhance Physical and Psychological Functioning Among Younger Women With Early-Stage Breast Cancer


ABSTRACT

Purpose
To identify factors that condition or moderate the impact of a previously described set of interventions on psychological and physical adjustment after diagnosis and treatment for early-stage breast cancer.

Patients and Methods
Younger women (age $\leq 51$ years, $N = 252$) with early-stage breast cancer within 2 months of having completed active nonhormonal adjuvant therapy were randomly assigned to a three-arm clinical trial, consisting of a control arm, an education arm, and a nutrition arm. Primary end points, assessed before random assignment and 4 and 13 months later, included mental functioning, physical functioning, and depressive symptoms. Four types of moderator variables were identified, including two sets reflecting psychosocial resources, specifically personality factors and factors related to the person’s social environment, a set reflecting demographic variables, and a set reflecting treatment and disease variables (including comorbidities).

Results
Psychosocial factors were more likely to moderate treatment effects than were demographic and disease-related factors, but the moderating effects of these psychosocial factors were limited to patients receiving the nutrition intervention. Patients with lower psychosocial resources benefited from the nutrition intervention, whereas patients with a greater amount of psychosocial resources did not.

Conclusion
Future trials of this type should stratify by or select for the moderating variables identified here (ie, dispositional pessimism, unmitigated communion, and negative social interaction) to establish more firmly their role in responses to psychosocial interventions. Effort should also be made to collect data to inform the delivery of interventions to those who might benefit the most.

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INTRODUCTION

A recently completed clinical trial, the Breast Cancer Recovery Project (BCRP), has shown that two brief psychosocial interventions, one involving a series of educational sessions and one involving a series of nutritional sessions, can reduce depressive symptomatology and enhance physical functioning among younger women diagnosed with early-stage breast cancer. This article examines whether the main intervention effects previously reported by Scheier et al were moderated or conditioned by other factors.

Although one must be cautious when interpreting moderation effects in randomized trials, it is useful to explore moderation effects because they can provide valuable insights about maximizing treatment outcomes. In this regard, past research on moderators suggests that interventions, particularly psychosocial interventions, may have the greatest impact on those persons who are at greatest risk for an adverse outcome. The risk might be a result of maladaptive traits that characterize the person’s personality, poor coping skills, or a weak social support network. Regardless of the source of the risk, patients with poorer psychosocial resources seem to benefit more from the interventions than patients with richer psychosocial resources, and we expected the same general pattern of results to emerge here.

Several different types of moderator variables were assessed. One set of variables included personality factors. Of particular interest were optimism and resilience, which have been consistently linked to better outcomes in previous research.
Primary outcomes included depressive symptoms (measured by an abbreviated form of the Center for Epidemiologic Studies Depression Scale36), mental health functioning, and physical health functioning (both quality-of-life variables measured by the Mental Component Summary score and Physical Component Summary score of the Medical Outcomes Study 36-Item Short Form Health Survey,29 respectively). Information about primary outcomes was collected at baseline (before random assignment) and 4 months and 13 months later (ie, immediately and 9 months after the intervention, respectively). Information about potential moderators of obtained treatment effects, which provides the focus of the present report, was collected at baseline.

Modrating Variables Assessed

Personality measures. Dispositional optimism was assessed by the Life Orientation Test–Revised30 (mean score = 16.73, standard deviation [SD] = 3.78, range = 1 to 24, α = .87; note that all of the descriptive statistics reported for all of the measures come from the present sample). Unmitigated communion was assessed by a nine-item scale created by Fritz and Helgeson31 (mean score = 29.72, SD = 5.20, range = 12 to 42, α = .75).

Social support/social environment measures. Four social environment variables were assessed. Negative social interaction, which refers to the extent to which the participant was receiving unhelpful support from her social environment, was assessed by five items from the Unsupportive Social Interactions Inventory32 (mean score = 7.92, SD = 2.39, range = 5 to 16, α = .61). A composite variable measuring partner support satisfaction was created by standardizing and then summing the scores on three separate scales, a 16-item relationship satisfaction measure1 a 16-item relationship conflict measure1 (both developed for this study), and a nine-item measure33 assessing the extent to which the partner was accepting of the participant’s surgical scars and not reacting in a negative emotional way (mean score = 0.00, SD = 2.62, range = −12.38 to 5.27, α = .82, for composite measure; average r among components of composite = 0.65). The quality of family interactions was assessed by standardizing and summing the scores on the Closeness, Conflict, and Expression subscales of the Moos and Moos34 Family Environment Scale (mean score = 0.80, SD = 2.47, range = −10.49 to 5.13, α = .77, for composite measure; average r among subscales = 0.52). Finally, perceived support was assessed using the six-item version of the Interpersonal Support Evaluation List (mean score = 21.67, SD = 2.95, range = 10 to 24, α = .85).35

Demographic measures. Demographic measures included age at the time of the baseline interview, marital status, and race/ethnicity. A measure of socioeconomic status was obtained by aggregating across standardized measures of years of education, occupation, and household income.

Disease- and treatment-related measures. Stage of cancer (either 0, I, or II) was used as a marker of disease severity. A treatment extensiveness index was created by standardizing and aggregating type of surgery (1 = lumpectomy, 2 = mastectomy) with type of adjuvant treatment received (0 = none, 1 = radiation or chemotherapy, 2 = both). Variables were standardized before they were summed. A change in menstrual status between diagnosis and baseline and weight gained or lost since diagnosis (both coded as yes or no) were combined to form an adverse effects of treatment index. In addition, respondents were given at baseline a checklist of potential comorbidities (eg, diabetes, migraines, arthritis, or angina) and were asked to check any conditions that they currently had. The number of conditions checked was used as a measure of comorbidities.

Statistical Analysis

Moderator variables that were categoric in nature were analyzed using standard analysis of covariance procedures, followed by analyses using simple contrasts. Moderator variables that were continuous in nature were analyzed using regression procedures outlined by Aiken and West.36 Variables were centered before they were entered into equations. Also following guidelines provided by Aiken and West,36 each moderator was run in two separate equations, one with the nutrition arm dummy coded and one with the education arm dummy coded. Those analyses revealing significant moderator effects were followed by analyses of simple slopes. To analyze the simple slopes, the relations between treatment arms and outcome were evaluated 1 SD above and below the mean of the moderator.

PATIENTS AND METHODS

Participants

The data for the present analyses were provided by the women (N = 252) who participated in the BCRP. The participants in the BCRP have been fully described elsewhere.1 Eligibility criteria included the following: stage 0, I, or II breast cancer; no history of treatment for other cancers; age younger than 51 years at the time of recruitment; within 2 months of completion of nonhormonal adjuvant therapy; and English speaking. The study was conducted in accordance with assurances approved by the Department of Health and Human Services and was approved by relevant institutional review boards. All participants signed informed consents.

Protocol

The protocol used in the BCRP has also been described fully elsewhere.1 In brief, participants in the BCRP trial were randomly assigned either to a control arm, which received standard medical care, or to one of two active treatment arms. Each active treatment arm received a series of four group sessions that met once a month for 4 consecutive months. Each monthly session lasted 2 hours. Participants in the education arm received information about their disease and treatment and were provided with a set of relevant coping skills, whereas participants in the nutrition arm received information on how to adopt and adhere to a low-fat, high-fruit, and high-vegetable diet.
For those categories of moderator variables in which more than one significant moderator emerged, a subsequent hierarchical regression analysis was conducted in which all significant moderators were entered simultaneously. These analyses also followed the data analytic strategy suggested by Aiken and West.36 Because we were interested in moderators of long-term intervention effects, the moderator analyses conducted were limited to treatment outcomes observed at the 9-month postintervention follow-up. Following Aiken and West,36 results for continuous moderator variables are reported as unstandardized regression coefficients. Finally, although the interventions did not affect mental health component scores in the original randomized clinical trial,3 analyses were conducted to test for moderators of the mental health component. None of the moderators studied here were significant. As a result, mental health functioning will not be discussed further.

**RESULTS**

**Depressive Symptoms**

Several variables moderated the impact of the nutrition intervention on depressive symptoms. With respect to social support/social environment variables, the effects of the nutrition intervention on depressive symptoms were moderated by perceptions of negative social interaction ($\beta = -0.59, SE = 0.26, \Delta R^2 = 0.02, P < .03$). Calculation of the simple slopes showed that the nutrition intervention significantly lowered depressive symptoms among participants who reported receiving high levels of negative social interaction ($\beta = -0.30, P < .001$). The effect of the nutrition intervention among participants who reported experiencing fewer negative social interactions was similar in form but much weaker and nonsignificant ($\beta = -0.05, P = .56$).

The effect of the nutrition intervention on depressive symptoms was also moderated by both personality characteristics, dispositional optimism ($\beta = 0.38, SE = 0.17, \Delta R^2 = 0.02, P < .03$) and unmitigated communion ($\beta = -0.24, SE = 0.12, \Delta R^2 = 0.01, P < .05$). With respect to dispositional optimism, analyses of simple slopes indicated that the nutrition intervention significantly reduced depressive symptoms among participants who were more pessimistic in outlook ($\beta = -0.31, P < .001$) but did not have a significant effect on depressive symptoms among those more optimistic in outlook ($\beta = -0.05, P = .55$). A conceptually similar pattern emerged for unmitigated communion, in that the nutrition intervention significantly lowered depressive symptoms among those with lower intrapersonal resources (ie, those higher in unmitigated communion; $\beta = -0.27, P < .01$) but did not affect depressive symptoms among those lower in unmitigated communion ($\beta = -0.05, P = .55$).

To determine the unique contribution of each of these three psychosocial moderator variables, a subsequent hierarchical multiple regression analysis was conducted in which all three moderators were entered simultaneously. Although the effect for negative social interaction approached significance ($P < .09$), none of the interaction terms attained statistical significance. Thus, it was apparently the shared variance among them that was producing the significant effects in the univariate analyses (even though the average absolute correlation among the moderators was only 0.25).

**Physical Functioning**

The only variable to moderate intervention effects on physical functioning was comorbidities. Number of comorbidities moderated the impact of both the nutrition intervention ($\beta = 0.83, SE = 0.36, \Delta R^2 = 0.02, P < .03$) and the education intervention ($\beta = -0.71, SE = 0.35, \Delta R^2 = 0.01, P < .05$). Analyses of simple slopes indicated that the nutrition intervention produced a significant improvement in physical functioning among those with more comorbidities ($\beta = -1.5, P < .05$). In contrast, the nutrition intervention had little impact among patients with fewer comorbidities ($\beta = -1.1, P > .18$). Comorbidities had the opposite effect on the education intervention. Those with fewer comorbidities were benefited by the education intervention ($\beta = .27, P < .001$), whereas those with more comorbidities were not ($\beta = .05, P > .49$).

Scheier et al1 have shown that two brief psychosocial interventions, one involving a series of educational sessions and one involving a series of nutritional sessions, can reduce depressive symptomatology and enhance physical functioning among younger women diagnosed with early-stage breast cancer. The present study compliments this earlier work by identifying who benefited most from those interventions. Several observations can be made with respect to the findings from this study.

First, some participants in the interventions benefited from them, and others did not. More specifically, the effect of the nutrition intervention on depressive symptoms was conditioned by optimism and unmitigated communion, as well as the magnitude of negative social interaction. In addition, comorbidities moderated the impact of both interventions on physical functioning.

Second, the moderation took a form that was generally consistent with prior expectations, in that the interventions were effective for those people who possessed more negative and maladaptive attributes, presumably because they have more to gain from what the interventions have to offer. In this respect, our findings are conceptually consistent with recent resource theories of stress and coping.37–42

The nature of the moderators that emerged in this study is noteworthy. That is, both dispositional optimism and social environment factors have been shown in previous studies3 to moderate the effects of psychosocial interventions in a manner that is conceptually consistent with the effects that were found here. Taken together, these studies suggest that these factors may play a particularly important role in deciding who might benefit from psychosocial interventions for breast cancer.

It is also important to note, however, that the nature of the interventions that are moderated varies somewhat across studies. Previous research1,3,5 has shown that psychosocial factors, such as the ones studied here, can moderate interventions that are similar to the one delivered in the education arm of the present study. Yet, in the present study, the effect of the education intervention was not conditioned by psychosocial factors. It is unclear why these differences across studies have emerged. Perhaps the differences between the present and prior research have to do with differences in the age of the samples studied (the present sample was younger) or in the time at which the interventions were targeted (at the end of nonhormonal adjuvant treatment in the present study v shortly after diagnosis and initiation of treatment in the previous studies). The attempt should be made in future research to better understand these differences in findings.

Why, in the present case, should psychosocial factors be important moderators of outcomes for women in the nutrition arm? One
possibility is that people with low psychosocial resources eat more poorly and so gain greater psychological benefit from a nutritional rather than an educational intervention. Alternatively, changes to one’s eating patterns may be more demanding or harder to enact than the changes called for in the educational intervention. People with low psychosocial resources may have been particularly bolstered by the structure and strategies that the nutrition intervention offered.

The findings with respect to comorbidities were also complex. Prior research has shown that those with more comorbidities may benefit more from interventions designed to reduce depressive symptomatology than those with fewer comorbidities. We did not find that comorbidities moderated the effects of our interventions on depressive symptoms. However, we did find that comorbidities moderated the effects of our interventions on physical functioning, but the results varied by intervention. With respect to the nutrition intervention, participants with more comorbidities (i.e., more burdened or greater need) benefited from the intervention, whereas those with fewer comorbidities did not. This effect is similar conceptually to the general pattern of effects obtained for our psychosocial moderators, as well as the results reported by Reynolds et al. However, the opposite was true when we examined the effects of the education intervention on individuals with different levels of comorbidities. In this arm, participants with fewer comorbidities benefited.

It is unclear why these differences emerge across outcomes and interventions, but the differences do suggest that the moderating effect of comorbidities is less straightforward than the moderating effects of factors that are more psychosocial in nature. One possibility involves the heterogeneous nature of comorbidities. Perhaps type of comorbidity interacts with the patient population under study, sometimes producing enhanced intervention effects and sometimes producing impaired intervention effects, depending on the specific patient population and the specific comorbidities that are assessed. Future research should be sensitive to this possibility.

Finally, it is of interest to note that the interventions were not moderated by any treatment or disease variable, including extensiveness of treatment and severity of disease. In this regard, it is noteworthy that conceptually similar results have recently been reported by others. The present findings thus add to the growing body of research that suggests that treatment and disease factors may not be as important as one might think in affecting psychosocial adjustment after breast cancer diagnosis and treatment.

Prior research has shown that education and nutrition interventions can be created that enhance psychological and physical functioning among younger women with early-stage breast cancer. The present findings add to our understanding by documenting that some women benefit from these interventions and others do not. In general, women benefited if they had poor or maladaptive intrapersonal and/or interpersonal resources. Future trials of this type should stratify by or select for these moderating variables to establish more firmly their role in responses to psychosocial interventions. The information provided by these moderation analyses will help inform the delivery of interventions to those who might benefit the most.
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