Promotion, Turnover and Compensation in the Executive Market

George-Levi Gayle  
_Carnegie Mellon University, ggayle@andrew.cmu.edu_

Limor Golan  
_Carnegie Mellon University, limor@andrew.cmu.edu_

Robert A. Miller  
_Carnegie Mellon University, ramiller@andrew.cmu.edu_

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Abstract

This paper is an empirical study of the market for managers, more specifically the effects of agency, human capital, and preferences on their promotion, tenure, turnover and compensation. From a large longitudinal data set compiled from observations on executives and their publicly listed firms, we construct a career hierarchy and report on its main features. Our summary results motivate a dynamic competitive equilibrium model, whose parameters we identify and estimate. Controlling for heterogeneity amongst firms, which differ by size and sector, and also managers, whose backgrounds vary by age, gender and education, our estimates are used to evaluate how important moral hazard and job experience are in jointly determining promotion rates, turnover and compensation.

1 Introduction

Chief executives are paid more than their subordinates, and internal promotions with the firm are positively correlated with wage growth.\(^1\) Since high ranking executives are almost always drawn from the lower ranks, usually from within the firm, it is tempting to conclude that part of the reward from working hard in a low rank is the chance of promotion to earn rents. Theory provides several possible explanations, ranging from human capital acquired on lower level job, to superior ability being revealed with experience leading to wage dispersion, or as the prize in a tournament played by lower ranked executives to induce hard work.\(^2\) The premise of all these explanations is the commonly held opinion that the CEO is better off than those he supervises. Yet several studies, conducted with data on executive compensation and returns from publicly traded firms, show quite conclusively that CEO compensation is more sensitive to the excess returns of firms than the

\(^{1}\)See Lazear (1992), Baker, Gibbs and Holmstrom (1994a), McCue (1996)

compensation of lower ranked executives.\textsuperscript{3} Thus at the upper levels of the career ladder, differently ranked jobs do not have the same characteristics. Whether one job is more desirable than another depends on the probability distribution of financial compensation that generates his income, as well as its nonpecuniary costs and benefits.

To the best of our knowledge, no one has attempted to quantify how much a CEO receives as a rent from human capital in management and leadership, and how much he is compensated for receiving a more volatile income. A small but growing literature on the structural estimation of moral hazard models investigates the empirical relationship between the principal’s return and the agent’s compensation, in order to quantify how incentives are used for inducing agents to work in the interests of their principals and truthfully revealing their hidden information.\textsuperscript{4} These studies find that estimates of the higher risk premium necessary to compensate a CEO for a more uncertain income relative to the second in command are of the same order of magnitude as differences in expected compensation. Such findings do not resonate with common opinion, because they imply the CEO receives very little pecuniary rent from his promotion to that position. Published work does not, however, integrate human capital and its behavioral consequences into an optimal contracting framework, confounding any attempt to gauge the degree of on-the-job training provided at lower ranks relative to the nonpecuniary value of holding a job at any given rank. More generally, the empirical importance of human capital in the executive labor market, and the role of promotions in this process, is unclear.\textsuperscript{5}

This paper is an empirical study of the effects of incentives, human capital, and preferences of managers, with goal of explaining the differences in the promotion, tenure, job turnover and compensation structure across managers using a dynamic competitive equilibrium model. Our data contain background information on executives, including age, gender, education, executive experience and the types of firms they work for, plus detailed information on their compensation and the financial returns of their firms and their rank within a career hierarchy. We identify and estimate a dynamic equilibrium model to analyze and disentangle the effects of competition in the market for managers using data on internal promotions, job turnover and the compensation of executives. estimate. Controlling for heterogeneity amongst firms and managers, our estimates are used to evaluate how important moral hazard and job experience are in jointly determining promotion rates, turnover and compensation.

The model is set up in the next two sections. Executives choose job, firm and effort level every period. They have preferences over jobs, particularly, effort is costly. These taste parameters vary across jobs and firms. In addition, every period managers privately observe a firm-job specific taste shock. The effort level is private information as well. While working they accumulate firm-specific and general human capital. We assume human capital accumulation on a job is greater when the manager exerts effort. The rate of human capital accumulation varies across jobs and firm as well, therefore, working in some firms and jobs may increase the manager’s stock of human capital. Firms offer contracts which provide incentives for managers to exert effort. Because exerting effort increases the

\textsuperscript{3}See Margiotta and Miller (2000) and Gayle and Miller (2008a, 2008b).
\textsuperscript{5}Frydman (2005) finds evidence on the increase importance of general skills in executive compensation.
manager’s stock of human capital, future promotion prospects provide incentives. Thus, variation in compensation across firms and jobs partially reflect the different opportunities to accumulate human capital and different promotion prospects. In addition, managers’ age and rank imply differences in career concerns affecting the optimal compensation schemes. The markets for executives is competitive. Managers have different stocks of human capital and compensation adjusts to clear the market for each skill set.

Identification of the parameters of the model is analyzed in Section 4, while our data is described in Section 5, where we define the job hierarchy and wage compensation. Our measure of compensation is comprehensive, and includes salary and bonus, stock and option grants, retirement benefits, as well as income directly attributable to holding securities in the firm in lieu of a widely diversified portfolio. The compensation data is augmented with data on the titles of the executives, along with their professional and demographic background compiled from the Marquis "Who’s Who". Compensation of the executives are sensitive to fluctuations in the abnormal returns. In fact, the firm’s excess return (over and above the market’s return) is the most important determinant of managerial compensation, suggesting the importance of incentives and moral hazard. We find that in fact the higher the executive’s rank in the firm, the more sensitive his compensation to the abnormal return. We also find that firm turnover is positively correlated with promotions and higher compensation.

Estimation is discussed in Section 6, while some preliminary estimates from the structural estimation are reported in the final section. We used four metrics to assess how much agency problems in executive markets are mitigated by their career concerns. Two of these measure the impact of an executive shirking rather than working, while the other two focus on the cost of eliminating the moral hazard problem. We find that firms are prepared to pay hardly anything to eliminate the moral hazard problem at the lower ranks, but that at the upper levels, the risk premium paid to executives for accepting an uncertain income stream that depends on the firm’s abnormal returns, are considerably greater. Career concerns greatly ameliorate the moral hazard problem for lower level executives, but their importance declines monotonically with promotion through the ranks. Overall our empirical findings, based on a large sample of executives employed by a broad cross section of publicly traded firms, demonstrate that the design of the hierarchy and the promotion process are important tools, used in conjunction with compensation schemes, for disciplining employees and aligning their interests to the goals of the organization.

2 The Model

Our model analyzes promotion, turnover and executive compensation, where expected value maximizing shareholders are subject to moral hazard from choices made by their risk averse expected utility maximizing executives, who are more informed than their employers about the value of their job matches. Executives earn returns from investment in human capital by gaining seniority within a position, from internal promotion, and turnover to other firms. These three factors, rooted in the technology of learning on the job, may induce them to trade off higher current income for better future prospects as their

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6 Gibbons and Murphy (1992) develop and empirically test a model of optimal contracts in the presence of career concerns in the market for CEOs.
career opportunities unfold. Behavior on the job is also affected by these three factors, as well as the compensation schedule, which depends on signals shareholders receive about managerial performance. Designed to align the goals of the firm with the executive, this variability induces a risk premium. We derive a competitive equilibrium where the optimal contracts of shareholders guide managerial decisions on job choice and effort on the job.

At the beginning of every period, equity returns of firms from decisions made in the previous period are revealed to everyone, the human capital state variables of executives are updated, and each executive is compensated by following the schedule of the previous period’s employment contract. Firms assess their demand for executives in the current period and advertise for executives internally and externally, by posting one-period contracts for positions within their firms. Then executives privately observe realizations of preference shocks and choose their consumption. They accept their most attractive employment offer, or quit management, and markets clear. Finally each executive chooses an effort level, a choice that is concealed from everyone else but nevertheless affect both his utility and the distribution of the returns of his firm realized at the beginning of the next period. Given the employment contracts offered by potential employers, executives sequentially maximize expected lifetime utility with respect to consumption, employment and effort level. This section develops the model.

2.1 Choices, Human Capital and Preferences

There are a finite number of firm types in the market indexed by $j \in \{1, ..., J\}$, with $j = 0$ representing retirement. There are $K$ different types of positions within each firm type $j$, indexed by $k \in \{1, ..., K\}$ and ranked in hierarchical order. We let $d_{jkt} \in \{0, 1\}$ indicate the manager’s job, his rank $k$ at firm $j$ in time period $t \in \{0, 1, \ldots\}$, and let $d_{0t}$ denote the indicator variable for retirement, which is an absorbing state. The $JK + 1$ choices are mutually exclusive, implying:

$$d_{0t} + \sum_{j=1}^{J} \sum_{k=1}^{K} d_{jkt} = 1$$

for all time periods $t \in \{0, 1, \ldots\}$ before retirement. Summarizing, $d_t \equiv (d_{0t}, d_{1t}, \ldots, d_{JKt})$ denotes the vector of job and rank choices an executive makes in period $t$. There are two activities within the firm, called working and shirking, denoted by $l_t \in \{0, 1\}$, where $l_t = 0$ means the manager shirks in period $t$ and $l_t = 1$ means the manager works. Only the manager observes his own effort.

Human capital, denoted by the vector $h_t$, is sequentially determined by the choices of the executive, it also include other executive’s characteristics such as age. Given his period $t$ choices of effort level $l_t$ in the $k^{th}$ rank at the $j^{th}$ firm, his human capital at the beginning of period $t + 1$ is determined by the mapping:

$$h_{t+1} \equiv H_{jk} (h_t) l_t + H'_{jk} (h_t) (1 - l_t)$$

where $h_0$ represents the initial endowment of the executive (such as fixed demographic characteristics such as gender and education). This specification encompasses three dimensions of how human capital is accumulated. The first relates to where it can be acquired, for example in lower ranks versus higher. The second dimension relates to where it might apply, such as to all firms versus only firms belonging to the same industry. Firm
specific experience at rank \( k \) for example, \( \sum_{s=0}^{t} d_{jkt-s} \), might increase productivity in firm \( j \) at rank \( k' \) more than elsewhere. The third dimension is who observes an executive’s human capital. In our model some attributes can be only directly observed by the executive, such as accumulated effort \( \sum_{k=1}^{K} \sum_{s=0}^{t} d_{jkt-s} l_{t-s} \), whereas other attributes are observed by everyone, such as total executive experience \( \sum_{k=1}^{K} \sum_{s=0}^{t} d_{jkt-s} \).

Executives are infinitely lived, and their preferences are characterized by the discounted sum of a time additively separable constant absolute risk aversion utility function, which is multiplicative in consumption and nonpecuniary factors. Human capital affects both the productivity of the firm, as discussed below, and also enter preferences directly, through the ease with which tasks are accomplished. Thus the preference parameters of a manager depend on his employer and rank \( (j, k) \), his human capital \( h_{t} \), and his effort level \( l_{t} \). We represent the preference parameters by the mapping \( \alpha_{jkt} (h_{t}) \) and normalize the utility parameters associated with retirement, by setting \( \alpha_{0kt} (h_{t}) \equiv 1 \) for all \( (k, l) \). For any job choice \( j \neq 0 \) we assume there is more disutility from working than shirking or, noting that exponential utility is negative, \( \alpha_{jkt} (h_{t}) > \alpha_{jk0} (h_{t}) \) for all \( h_{t} \). An individual taste shock indexed by firm, position and time, also affects current utility. Denote this shock by \( \varepsilon_{jkt} \) if workplace position \( (j, k) \) is selected, and by \( \varepsilon_{0t} \) if the executive retires. For notational convenience we assume, that if the executive retires in period \( t \), then \( \varepsilon_{0s} \equiv 0 \) for all \( s > t \). Thus life-time utility can be summarized as:

\[
- \sum_{t=1}^{\infty} \beta^{t} \text{exp} (-\rho c_{t}) \left\{ d_{0t} \text{exp} (-\varepsilon_{0t}) + \sum_{j=1}^{J} \sum_{k=1}^{K} d_{jkt} [\alpha_{jk0} (h_{t}) (1 - l_{t}) + \alpha_{jkt} (h_{t}) l_{t}] \text{exp} (-\varepsilon_{jkt}) \right\}
\]

where \( \beta \) is the subjective discount factor, \( \rho \) is the constant absolute risk aversion parameter, and if \( d_{0t} = 1 \) then \( d_{0s} = 1 \) for all \( s > t \).

We assume there exists a complete set of markets for all publicly disclosed events relating to commodities, with price measure \( \Lambda_{t} \) and derivative \( \lambda_{t} \). This implies that consumption by the manager is limited by a lifetime budget constraint, which reflects the opportunities he faces as a trader and the expectations he has about his compensation. The lifetime wealth constraint is endogenously determined by the manager’s work activities. By assuming markets exist for consumption contingent on any public event, we effectively attribute all deviations from the law of one price to the particular market imperfections under consideration. Let \( e_{t} \) denote the endowment at date \( t \). We also measure \( w_{jkt},t+1 \), the manager’s compensation for employment in position \( k \) at firm type \( j \) at the beginning of period \( t+1 \), in units of current consumption. To indicate the dependence of the consumption possibility set on the set of contingent plans determining labor supply and effort, we define \( E_{t} [\bullet | l_{t}, d_{t}, h_{t}] \) as the expectations operator conditional on work and effort level choices at time \( t \), the subscript on the operator indicating shocks in the commodities market. The budget constraint can then be expressed as:

\[
E_{t} [\lambda_{t+1} e_{t+1}| l_{t}, d_{jkt}, h_{t}] + \lambda_{t} c_{t} \leq \lambda_{t} e_{t} + E_{t} [\lambda_{t+1} w_{jkt+1}| l_{t}, d_{jkt}, h_{t}]
\]  

(2)
2.2 Firms

We assume that the value of executive work to the firm is additive, an assumption of convenience that suppresses the role of teamwork and organizational capital. Specifically we assume that an executive with background \( h_k \) in the \( k^{th} \) rank contributes \( F_{jk}(h_k) \) to the the \( j^{th} \) firm in period \( t \). Let \( e_{jt} \) denote the equity value of firm \( j \) at time \( t \). Letting \( \pi_{t+1} \) denote the return on the market portfolio, we can now express the \( j^{th} \) firm’s value at the beginning of period \( t+1 \) as:

\[
e_{jt}(\pi_{t+1} + \pi_{j,t+1}) + \sum_{k=1}^{K_j} [F_{jk}(h_k) - w_{jk,t+1}]
\]

where \( \pi_{j,t+1} \) is a random variable drawn from a probability distribution the probability density function, which depends on the characteristics of the firm and managerial effort. We interpret \( \pi_{j,t+1} \) as the abnormal return to the firm before factoring in the skills of the management team and their aggregate compensation.

Moral hazard arises in this model because the probability distribution for \( \pi_{j,t+1} \) depends on the effort of all the executives. Let \( f_j(\pi_{j,t+1}) \) denote the probability density function for \( \pi_{j,t+1} \) conditional on all the executives working diligently, and let \( f_j(\pi_{j,t+1}) g_{jk}(\pi|h_t) \) denote its probability density function when all except the \( k^{th} \) ranked executive with human capital \( h_t \) working diligently, where \( g_{jk}(\pi_{j,t+1}|h_t) \) is a bounded strictly positive continuous function with \( E[g_{jk}(\pi_{j,t+1}|h_t)] = 1 \). A conflict of interest between owners and managers arises because managers prefer to shirk rather than work, meaning \( \alpha_{1jk}(h_t) > \alpha_{0jk}(h_t) \), but the expected return to firms is increasing in the number of its executives who work, meaning that for all \( h_t \):

\[
\int \pi f_j(\pi) d\pi > \int \pi f_j(\pi) g_{jk}(\pi|h_t) d\pi
\]

We assume that firms only write contracts that motivate the executive to work diligently, an assumption that is notationally cumbersome but conceptually easy to relax. In a more general setting firms would offer both types of contract, potentially doubling the number of equilibrium job offers available to each executive. A compensating differential would be offered for following the firm’s goals rather than shirking on the job. Moreover compensation from a shirking contract would not depend on the firm’s profitability, and hence the executives who accepted such offers would not be paid a risk premium. Consequently executives who accepted shirking contracts would earn much less than the top executives on average, and may not even hold the same title. For these reasons we consolidated the activity of equilibrium shirking on the job within the position of retirement.

3 Individual Optimization

The executive chooses his consumption stream as his income accrues from a sequence of lotteries with prizes of \( w_{jk,t+1} \) until he retires, and also as he receives nonpecuniary (scaled utility) benefits from participating of \( \alpha_{jk}(h_t) \exp(-e_{jk,t}^*) \), where \( e_{jk,t}^* \) is the value of the period \( t \) disturbance when lottery \( (j, k) \) is selected in period \( t \). We first solve for the stochastic sequence of executive’s consumption and savings choices as a function of
the compensation schemes offered by different firms given the career choices he makes when he always works diligently. Then we solve for the their workplace and rank choices, conditional on working diligently each period, a dynamic discrete choice problem. Finally we analyze the incentives that determine effort level, and what restrictions must be placed on contracts to induce executives prefer working diligently rather than shirking along the equilibrium path.

3.1 Consumption and Saving

Suppose, for the moment, that an executive always works diligently, and ignores the job choices of the other executives in the firm. Then we can denote by \( p_{jkt} (h_t) \) the probability that the manager optimally selects job \((j, k)\) in period \(t\) given characteristics \(h\), found by integrating \( d_{jkt} \) over \( (\varepsilon_{0t}, \varepsilon_{1t}, \ldots, \varepsilon_{Jkt}) \). In this section we assume that the executive can work at most a finite number of periods before retiring, denoted by \(T < \infty\), but we relax this restriction in our summary of the competitive equilibrium concluding the next section.

In general the indirect utility function depends on all the state contingent prices, but for the exponential utility specialization, just two securities suffice. Let \( b_t \) denote the period \(t\) price of a infinitely lived bond, let \( a_t \) denote the price of a security that pays off the (random) dividend \((\ln \lambda_s - s \ln \beta - \ln \Lambda)\) is period \(s\), and define:

\[
V_{jk, t+1} = \exp (-\rho w_{jk, t+1}/b_{t+1})
\]

as the risk adjusted utility weight for receiving compensation \(w_{jk, t+1}\) at the beginning of period \(t+1\) for working \((j, k)\) in period \(t\). For all \(s \in \{1, \ldots, T\}\) we set \(A_0 (h_t, b_t) \equiv 1\) and recursively define \(A_s (h_t)\) as:

\[
A_s (h_t, b_t) = p_{0t} (h_t) E \left[ \exp (-\varepsilon_{0t}^* / b_t) \right] + \sum_{j=1}^{J} \sum_{k=1}^{K} \left\{ p_{jkt} (h_t) [\alpha_{1jk} (h_t)]^{1/\beta_t} E \left[ \exp \left( -\varepsilon_{jkt}^* / b_t \right) \right] \right\} (3)
\]

\[
\times E_t [v_{jk, t+1} A_{s-1} (H_{jk} (h_t), b_{t+1})]^{1-1/\beta_t} (4)
\]

Letting \(c_t^o\) denote the optimal consumption in period \(t\), we now define the value function as:

\[
V_t (h_t, a_t, b_t) = -E_t \left\{ \sum_{s=t+1}^{\infty} \sum_{j=1}^{J} \sum_{k=1}^{K} \beta^{s-t} \exp (-\rho c_t^o) \left[ p_{0s} (h_s) \exp (-\varepsilon_{0s}) + p_{jks} (h_s) \alpha_{1jk} (h_s) \exp (-\varepsilon_{jks}) \right] \right\}
\]

The next lemma shows that \(c_t^o\), the optimal consumption rule for the manager, and the value function \(V_t (h_t, a_t, b_t)\) is formed from \(A_s (h_t)\) as Lemma 2 shows that both are mappings of \(A_s (h_t, b_t)\).

Lemma 1

\[
V_s (h_t, a_t, b_t) = -A_s (h_t, b_t) b_t \exp \left( -\frac{a_t + \rho c_t}{b_t} \right) c_t^o = \frac{e_t}{b_t} + \frac{t}{\rho} \ln \beta + \frac{a_t}{\rho b_t} - \left( \frac{b_t - 1}{\rho b_t} \right) e_{jkt}^* \ln \alpha_{1jk} (h_t) - \rho^{-1} \ln E_t [v_{jk, t+1} A_{s-1} (H_{jk} (h_t), b_{t+1})]
\]
Noting that \( b_t \exp [- (a_t + \rho e_t) / b_t] \) is the well known formula for the valuation function associated with exponential utility, we interpret \( A_s (h_t, b_t) \), the weight on \( b_t \exp [- (a_t + \rho e_t) / b_t] \), as a weight that values the executives career prospects given his human capital \( h_t \) and aggregate conditions \( (a_t, b_t) \) and future career of length \( s \). By inspection the index \( A_s (h_t, b_t) \) takes only strictly positive values. From the formula for \( V_s (h_t, a_t, b_t) \), lower values of \( A_s (h_t, b_t) \) are associated with a higher investment value and a higher valuation function. Finally note that if \( d_{0t} = 1 \) and the executive retires, then \( p_{0t} (h_t) = 1 \) and \( \varepsilon_{0t} = 0 \) for all periods \( \tau > t \), implying \( A_s (h_\tau, b_\tau) = 1 \).

The first three terms of the formula for optimal consumption are familiar, spending the interest on the endowment, \( e_t b_t^{-1} \), discounting consumption over time due to impatience, \( \rho^{-1} t \log \beta \), and adjusting for aggregate risk, \( a_t (p b_t)^{-1} \). The next term depends on the effects of nonpecuniary features of the job on the marginal utility of consumption, that which is publicly observed, \( \alpha_{1jk} (h_t) \), and the hidden component \( \varepsilon_{jkt} \). Since both components are specific to rank and firm, so is optimal consumption. Finally consumption in the current period also depends on \( E_t [v_{jk, t+1} A_{s-1} (H_{jk} (h_t), b_{t+1})] \), where \( v_{jk, t+1} \) is compensation due next period and \( A_{s-1} (H_{jk} (h_t), b_{t+1}) \) is an index of the investment value of the job opportunities next period reflecting growth in wealth from future work.

### 3.2 Job Choices

The supply of executives is determined by the job choices they make over their career. Given the optimal rule for consumption, the optimal position is found by selecting rank \( k \) in firm \( j \) at time \( t \) given human capital \( h_t \) and private value \( \varepsilon_{jkt} \) to sequentially maximize the sum of current utility, that is \( \alpha_{1jk} (h_t) \exp (\rho c_t^3 - \varepsilon_{jkt}) \) or \( \exp (\rho c_t^3 - \varepsilon_{0t}) \) in the case of retirement, plus the one period discounted expected value of future optimized utility, \( \beta E_t [V_{t+1} (H_{jk} (h_t), a_{t+1}, b_{t+1})] \) or \( \beta E_t [V_{t+1} (H_0 (h_t), a_{t+1}, b_{t+1})] \). Substituting in the formulas for optimal consumption and the valuation function reduce the job choice problem to the following formulation.

**Lemma 2** If \( d_{0s} = 0 \) for all \( s \in \{0, \ldots, t - 1\} \) and any \( t \in \{0, \ldots, T\} \), then the optimal job choice indicators \( d_t \) are picked to minimize:

\[
\varepsilon_{0t} d_{0t} + \sum_{j=1}^{J} \sum_{k=1}^{K} d_{jkt} \{ \varepsilon_{jkt} - \ln \alpha_{1jk} (h_t) - (b_t - 1) \ln E_t [v_{jk, t+1} A_{s-1} (H_{jk} (h_t), b_{t+1})] \}
\]

For example if \( \varepsilon_{jkt} \) is identically and independently distributed Type I Extreme Value, then the recursion for \( A_s (h_t) \) and the choice probabilities simplify tot the formulas given in following lemma.

**Lemma 3** If \( \varepsilon_{jkt} \) is independently and identically distributed as extreme value Type I with location and scale parameters \((0, 1)\), then:

\[
E [\exp (\varepsilon_{jkt} / b_t)] = p_{jkt}^{-1 / b_t} \Gamma [(b_t - 1) / b_t]
\]

\[
A_s (h_t, b_t) = p_{0t} (h_t)^{1 - \frac{1}{b_t}} \Gamma [(b_t - 1) / b_t]
\]

\[
+ \sum_{j=1}^{J} \sum_{k=1}^{K} p_{jkt} (h_t)^{1 - \frac{1}{b_t}} \alpha_{1jk} (h_t) E_t [v_{jk, t+1} A_{s-1} (H_{jk} (h_t), b_{t+1})]^{1 - \frac{1}{b_t}} \Gamma [(b_t - 1) / b_t]
\]
and:

\[
p_{jkt}(h_t) = \frac{\alpha_{jkt}(h_t) E_t[v_{jk,t+1}^{*} A_{s-1}(H_{jk}(h_t), b_{t+1})]^{b_t}}{1 + \sum_{j'} \sum_{k'} \alpha_{j'k'}(h_t) E_t[v_{j'k',t+1}^{*} A_{s-1}(H_{j'k'}(h_t), b_{t+1})]^{b_t}}
\]  

(5)

### 3.3 Work Effort

In our framework the support of \( \pi_{jkt} \) does not depend on the effort choice, and executives work diligently in equilibrium. Consequently the action of shirking is not detected by shareholders, since all shirking outcomes that shareholders observe can be rationalized by executive working hard. Similarly, since \( \varepsilon_{jkt} \) is private information and has full support, job choices made at the beginning of each period can be rationalized by a history of always working diligently. So regardless of what the manager chooses, and more generally what outcomes shareholders observe, they update their beliefs of his human capital, which we denote by \( h_{t}' \), as if he never strayed from the equilibrium path. Therefore the law of motion of \( h_{t}' \) is given by \( h_{t+1}' = H_{jk}(h_t') \).

Consider an executive who, after deviating from equilibrium, accumulated human capital of \( h_t \) when shareholders believe he has \( h_{t}' \). His conditional choice probabilities now depend on both \( h_t \) and \( h_{t}' \), because he knows he has \( h_t \) human capital, which stochastically determines his true productivity, yet he is paid as if he has \( h_{t}' \). We denote them by \( p_{jkt}(h_t, h_t') \). Given job choice \( (j, k) \), the executive’s state variables are then updated using the formula \( h_{t+1} = H_{jk}(h_t) \) if he works diligently and \( H_{jk}'(h_t) \) if he shirks. Analogously to the definition of \( A_s(h_t) \) we define the recursion

\[
B_s(h_t, h_t', b_t) = p_{0t}(h_t, h_t') E\left[\exp(-\varepsilon_{0t}'/b_t)\right] + \sum_{(j,k)} \left[ p_{jkt}(h_t, h_t') E\left[\exp(-\varepsilon_{jkt}'/b_t)\right] \times \max\left\{ \frac{\alpha_{1jkt}(h_t)^{1/\tau} E_t[v_{jk,t+1}^{*} B_s(h_t, b_{t+1})]^{1-\frac{1}{\tau}}}{\alpha_{0jkt}(h_t)^{1/\tau} E_t[v_{jk,t+1}^{*} (\pi_{j,k,t+1}, h_t) B_{s-1}(H_{jk}(h_t), b_{t+1})]^{1-\frac{1}{\tau}}}, \frac{\alpha_{0jkt}(h_t)}{\alpha_{1jkt}(h_t)}\right\}\right]
\]

This recursion yields the following generalization to the results we derived for equilibrium behavior, determining optimal consumption, job choice and effort selection off the equilibrium path.

**Lemma 4** For all \((t, j, k, h_t, h_t')\) define the indicator variable \( l_{jkt} = l_{jkt}^{b_t}(h_t, h_t') \) by setting \( l_{jkt} = 1 \) if the inequality:

\[
\frac{E_t[v_{jk,t+1}^{*} B_{s-1}(H_{jk}(h_t), h_t', b_{t+1})]}{E_t[v_{jk,t+1}^{*} g(\pi, h_t) B_{s-1}(H_{jk}(h_t), h_t', b_{t+1})]} \leq \left[\frac{\alpha_{0jkt}(h_t)}{\alpha_{1jkt}(h_t)}\right]^{1/(b_t-1)}
\]

is satisfied, with \( l_{jkt} = 0 \) otherwise. Then the executive optimally selects his position and firm by choosing \( d_t \) to minimize:

\[
\varepsilon_0 d_{0t} + \sum_{j=1}^{J} \sum_{k=1}^{K} d_{jkt} \left\{ \varepsilon_{jkt} - \alpha_{1jkt}(h_t) - (b_t - 1) \ln \frac{E_t[v_{jk,t+1}^{*} B_{s-1}(H_{jk}(h_t), h_t', b_{t+1})]}{E_t[v_{jk,t+1}^{*} g(\pi_{j,k,t+1}, h_t) B_{s-1}(H_{jk}'(h_t), H_{jk}(h_t), b_{t+1})]} \right\}
\]

\[
- \left(1 - l_{jkt}\right) \left[\ln \frac{\alpha_{0jkt}(h_t)}{\alpha_{1jkt}(h_t)} - (b_t - 1) \ln \frac{E_t[v_{jk,t+1}^{*} g(\pi_{j,k,t+1}, h_t) B_{s-1}(H_{jk}'(h_t), H_{jk}(h_t), b_{t+1})]}{E_t[v_{jk,t+1}^{*} B_{s-1}(H_{jk}(h_t), h_t', b_{t+1})]}\right]
\]

9
Optimal effort $l^o_t$ and consumption $c^o_t$ are respectively determined from the equations:

$$l^o_t = \sum_{j=1}^{J} \sum_{k=1}^{K_j} d_{jkt}^o l^o_{jkt}$$

and:

$$c^o_t = \frac{e_t}{b_t} + \frac{t}{\rho} \log \beta + \frac{\alpha_t}{\rho b_t} - \left( \frac{b_t - 1}{\rho b_t} \right) \left[ d^o_{jkt} x^o_{jkt} + \sum_{(j,k)} d^o_{jkt} x^o_{jkt} \ln \alpha_{1jk} (h_t) \right]$$

$$-\rho^{-1} \sum_{(j,k)} d^o_{jkt} \ln \left( E_t \left[ v_{jkt, t+1} B_{s-1} \left( l^o_{jkt} H_{jkt} (h_t) + (1 - l^o_{jkt}) H'_{jkt} (h_t), H_{jkt} (h'_t), b_{t+1} \right) \right] \right)$$

Whether the executive shirks or not depends on the relative benefits to his current utility, $\alpha_{1jk} (h_t) - \alpha_{0jk} (h_t)$, how it affects expected lifetime utility through compensation, $E_t[\{v_{jkt, t+1} (\pi) [1 - g (\pi, h_t)] | h_t, h'_t\}]$, and the differential investment value from working diligently versus shirking. Human capital is associated with higher compensation and nicer work conditions in the future, so typically:

$$B_{s-1} \left( H'_{jkt} (h_t), H_{jkt} (h'_t), b_{t+1} \right) > B_{s-1} \left( H_{jkt} (h_t), H_{jkt} (h'_t), b_{t+1} \right)$$

In lower level jobs held at the beginning of a career, the differential investment value might be such an important force motivating executives that:

$$\frac{E_t \left\{ B_{s-1} \left[ H_{jkt} (h_t), H_{jkt} (h'_t), b_{t+1} \right] \right\}}{E_t \left\{ B_{s-1} \left[ H'_{jkt} (h_t), H_{jkt} (h'_t), b_{t+1} \right] \right\}} \leq \frac{\alpha_{0jk} (h_t)}{\alpha_{1jk} (h_t)}^{1/(b_t - 1)}$$

obviating the need to tie remuneration to the abnormal returns of the firm.

### 4 Competitive Equilibrium

In our model, shareholders assess their demand for executives in the current period and post one-period contracts for positions within their firms to maximize their expected return from executive employment subject to two constraints. Firms must achieve a target acceptance rate for all positions and all executive backgrounds in the competitive equilibrium, thus satisfying a participation constraint. Every contract also satisfies an incentive compatibility constraint that induces each executive to work diligently in equilibrium. In competitive equilibrium each firm (type) $j \in \{1, \ldots, J\}$ fills its positions $k \in \{1, \ldots, K_j\}$ with managers with skill types $h_k$ and at compensation levels $w_{jkt, t+1} (h_k)$ that are at least as profitable for the firm as any alternative skill set of their management team, and every executive chooses his most desirable position and work effort given the menu of contracts offered by all firms to those endowed with his skills. Entry into the market for executives drives the rents from contracting with them to zero. This section derives the optimal one period contract, and proves the existence of a unique competitive equilibrium.
4.1 Optimal Contracting

The demand for executives reflects their potential to add value to the firm. We express the demand by the \( j^{th} \) firm for an executive with human capital of \( h_t \) to fill the \( k^{th} \) position at time \( t \) as a well defined probability, denoted by \( P_{jkt} (h_t) \). Later in this section we derive this demand probability as a function of the model’s primitives in competitive equilibrium. To achieve a success rate of \( P_{jkt} (h_t) \), the firm must offer a sufficiently attractive compensation package to elicit this supply, setting \( P_{jkt} (h_t) \leq p_{jkt} (h_t) \). Substituting the expressions we obtained from the supply side discrete choice problem:

\[
P_{jkt} (h_t) \leq \Pr \left\{ (j, k) \in \arg \min \left[ \varepsilon_{0t}, \ v_{j'k't} - \ln \alpha_{1j'k'} (h_t) - (b_t - 1) \ln E_t \{ v_{j'k', t+1} A_{s=1} [H_{j'k'} (h_t), b_{t+1}] \} \right] \right\}
\]

Denoting the conditional demand probability simplex by \( P_t (h_t) \equiv \{ P_{11t} (h_t), \ldots, P_{JKt} (h_t) \} \), where the probability of retirement is simply \( 1 - \sum_{(j,k)} P_{jkt} (h_t) \), and noting that the logarithm of utility from retirement is \( \varepsilon_{0t} \), by Proposition 1 of Hotz and Miller (1993), there exists a mapping \( q (P) \) from the simplex to \( R^{JK} \) such that this inequality is met if and only if:

\[
q_{jk} [P_t (h_t)] \leq - \ln \alpha_{1jk} (h_t) - (b_t - 1) \ln E_t \{ v_{jk, t+1} A_{s=1} [H_{jk} (h_t), b_{t+1}] \}
\]

The left side of the inequality, \( q_{jk} [P_t (h_t)] \), is the expected value of the disturbance, \( \varepsilon_{jkt}^* \), for an executive who is on the cusp of accepting the \( (j,k) \) position over his alternative set when, given the conditions of the job, summarized by compensation \( v_{jk, t+1} \), investment value \( A_{s=1} [H_{jk} (h_t), b_{t+1}] \), and known nonpecuniary benefits \( \alpha_{1jk} (h_t) \), the \( j^{th} \) firm will fill the position with probability \( P_{jkt} (h_t) \). When \( \varepsilon_{jkt} \) is Type 1 Extreme value it is well known that \( q_{jk} (P_t) = - \ln (P_{jkt}) \), so in this case the participation constraint reduces to:

\[
P_{jkt} (h_t) \geq \alpha_{1jk} (h_t) E_t \{ v_{jk, t+1} A_{s=1} [H_{jk} (h_t), b_{t+1}] \}^{(b_t - 1)}
\]

Raising compensation, such as increasing \( v_{jk, t+1} \) by a positive constant, better working conditions, represented by lower values of \( \alpha_{1jk} (h_t) \), and higher investment values, that is lower values of \( A_{s=1} [H_{jk} (h_t), b_{t+1}] \), reduce the right side of this inequality and thus help the firm attain any target demand probability \( P_{jkt} (h_t) \).

Along the equilibrium path \( h_t = h'_t \), and \( l_{jkt} = 1 \) for all \( j, k \), implying \( A_s (h_t, b_t) = B_s (h_t, h_t, b_t) \) for all \( s \). Thus to maintain incentives along the equilibrium path the inequality in Lemma 3 simplifies to:

\[
\alpha_{1jk} (h_t)^{1/(b_t - 1)} E_t \{ v_{jk, t+1} A_{s=1} [H_{jk} (h_t), b_{t+1}] \} \leq \alpha_{0jk} (h_t)^{1/(b_t - 1)} E_t \{ v_{jk, t+1} g (\pi, h_t) B_{s=1} [H_{jk} (h_t), H_{jk} (h_t), b_{t+1}] \} \equiv U (h_t)
\]

Both constraints can be expressed as linear in \( v_{jk, t+1} \), and the objective function, the expected wage bill \( E_t (w_{jk, t+1}) \) can be expressed as a concave function of \( v_{jk, t+1} \), namely \( E_t (\ln v_{jk, t+1}) \). Maximizing \( E_t (\ln v_{jk, t+1}) \) subject to the two constraints then yields the cost minimizing contract.
Lemma 5 The cost minimizing contract that elicits high effort and attracts a manager with experience \( h_t \) to the \( k \)th position in the \( j \)th firm at time \( t \) with probability \( P_t (h_t) \) is:

\[
w_{jk,t+1} (\pi, h_t) = \frac{b_{t+1}}{\rho} \left\{ \frac{1}{(b_t - 1)} \ln \alpha_{1jk} (h_t) + \ln A_{s-1} [H_{jk} (h_t), b_{t+1}] + q_{jk} [P_t (h_t)] + \ln \left[ 1 - \eta g_{jk} (\pi, h_t) + \eta \left[ \frac{\alpha_{1jk} (h_t)}{\alpha_{0jk} (h_t)} \right]^{1/(b_t - 1)} A_{s-1} [H_{jk} (h_t), b_{t+1}] / B_{s-1} (H_{jk} (h_t), H_{jk} (h_t), b_{t+1}) \right] \right\}
\]

where \( \eta \) is the unique positive root to

\[
\int \left[ \frac{\eta^{-1} f_j (\pi) \left[ \frac{\alpha_{1jk} (h_t)}{\alpha_{0jk} (h_t)} \right]^{1/(b_t - 1)} A_{s-1} [H_{jk} (h_t), b_{t+1}] / B_{s-1} (H_{jk} (h_t), H_{jk} (h_t), b_{t+1}) - g_{jk} (\pi, h_t) \right] d\pi = 1
\]

Equation 6, the compensation schedule characterizing the optimal contract, decomposes into four additive pieces. The first term, \( b_{t+1} \log \alpha_{1jk} (h_t) / \rho (b_t - 1) \), is the amount that leaves a manager indifferent between retiring and accepting position \((j, k)\) if the private values are the same across all the choices, there is no investment value from accepting the position, and the compensation is fixed. The second term, \( b_{t+1} \log [A_{s-1} (H_{jk} (h_t), b_{t+1}) / \rho \), is the investment value from the position, a wage discount that offsets higher future expected earnings. The next term, \( b_{t+1} g_{jk} [P_t (h_t)] / \rho \), sets compensation levels to make the position sufficiently attractive to the executive in the \( P_{jk} (h_t) \) fractile. The expected value of the last term is a risk premium for taking a position whose compensation depends on the firm’s financial returns, and is therefore uncertain.

When human capital accumulation does not depend on effort so all human capital is public information, then \( h_t = h_t^t \) at every outcome. Since \( A_s (h_t, b_t) = B_s (h_t, h_t, b_t) \) the incentive compatibility constraint in this case reduces to the static model:

\[
E[v_{jk,t+1} (\pi) g (\pi, h_t) | h_t] \leq \left[ \frac{\alpha_{1jk} (h_t)}{\alpha_{0jk} (h_t)} \right]^{1/(b_t - 1)} E[v_{jk,t+1} (\pi) | h_t]
\]

In this case the \( A_{s-1} [H_{jk} (h_t), b_{t+1}] \) and \( B_{s-1} [H_{jk} (h_t), H_{jk} (h_t), b_{t+1}] \) terms cancel each other in (6) and (7). Consequently the only difference in the optimal contract distinguishing a model with human capital that depends on past job choices, from a model of pure moral hazard without any human capital, is the investment cost component of compensation \( b_{t+1} \log [A_{s-1} (H_{jk} (h_t), b_{t+1}) / \rho (b_t - 1) \). Regardless of whether human capital is observed or not, jobs associated with promotion prospects to higher paid jobs command an offsetting negative compensating differential, reflected in lower values of \( A_{s-1} (H_{jk} (h_t), b_{t+1}) \). But when human capital is not observed, the prospect of promotion also ameliorates incentive problems, and would predict that lower level jobs on fast promotion tracks require less incentive pay.

Note that equilibrium compensation depends on the position within the firm (since tasks vary with the position), the applicant’s known characteristics (since managers with different backgrounds bring different skills to the firm), the firm’s random return next period (since this is a signal shareholders receive about unobserved managerial effort), but not the characteristics of the other executives in the management team (since, by our separability assumption, the manager’s value to the firm is independent of the composition
of the other executives on the management team). In our framework we assume firms cannot commit to long term multiperiod contracts with their executive staff. If human capital depends on employment history but not on effort, then the optimal long term contract decentralizes to a sequence of short term contracts, obviating the need to consider anything but one period contracts. However if human capital is a function of unobserved effort, then there are benefits to shareholders from committing. In that scenario, the optimal long term compensation contract takes into account the signals a firm receives from abnormal returns about previous firm specific unobserved investments in human capital made by its workers. Absent a commitment device, firms engage in sequentially optimal short term one period contracts of this type.

4.2 Equilibrium Job Assignment

The demand for executive services is determined by a zero profit condition imposed in equilibrium:

\[
\int w_{jk,t+1} (\pi, h) f_j (\pi) d\pi = F_{jk} (h)
\]

Note that we are not assuming shareholders are unable to extract rent from hiring managers, merely that the expected surplus has already been impounded into the value of the firm before any contracts are written. Define \( W_{jk} (h_t, b_t) \) as compensation paid to an executive less the compensating differential associated with the unobserved component:

\[
W_{jks} (h_t, b_t) = \frac{b_{t+1}}{\rho} \left\{ \frac{(b_t - 1)^{-1} \ln \alpha_{jk} (h_t) + \ln [A_{s-1} (H_{jk} (h_t), b_{t+1})]}{1 - \eta g_{jk} (\pi, h_t) + \eta \frac{A_{s-1}[H_{jk}(h_t),b_{t+1}]}{B_{s-1}[H'_{jk}(h_t),H_{jk}(h_t),b_{t+1}]} b_{t+1}} \right\}
\]

Then from the optimal contract given in Equation (6):

\[
E_t [w_{jk,t+1} (\pi, h_t)] = E_t [W_{jk} (h_t, b_t) + \rho^{-1} b_{t+1} q_{jk} (P_t (h_t))] = F_{jk} (h_t)
\]

Inverting we obtain in recursive form

\[
P_{jk} (h_t) = q_{jk}^{-1} \left\{ \rho \left[ \frac{F (h_t) - E_t [W (h_t, b_t)]}{E_t [b_{t+1}]} \right]\right\}
\]

where \( F (h_t) - E_t [W (h_t, b_t)] \) is the vector formed from arraying \( F_{jk} (h_t) - W_{jk} (h_t, b_t) \). Using a backwards induction procedure we recursively solve out to get \( p_{jk} (h) \). In the Type 1 Extreme Value specialization the support of \( \varepsilon_{jk} \) is the real line, so the supply probabilities \( p_{jk} (h) \) are strictly positive for all \((h,j,k,t)\) which in turn implies \( p_{jk} (h) \) is strictly positive too:

\[
P_{jk} (h_t) = \exp \left\{ \rho \left[ \frac{W (h_t, b_t) - F (h_t)}{E_t [b_{t+1}]} \right]\right\}
\]

Upon recursively solving for \( P_{jk} (h_t) \) we can in principle express the contract purely as a mapping of the primitives of the model by sequentially recovering \( A_{s-1} [H_{jk} (h_t), b_{t+1}] \) and \( B_{s-1} [H'_{jk}(h_t),H_{jk}(h_t),b_{t+1}] \).

\(^7\) The proof of this statement follows arguments developed in Fudenberg, Holmstrom and Milgrom (1990).
4.3 Extension to Infinite Horizon

The final lemma of our theoretical analysis extends our analysis to the infinite horizon case \( T = \infty \).

**Lemma 6** Suppose \((h_t, a_t, b_t)\) is a Markov process and, conditional on \((h_t, a_t, b_t)\), \(\varepsilon_{jkt}^*\) is identically and independently distributed. Then there exists an \((h_t, h'_t, b_t)\) which uniquely satisfies the fixed point:

\[
B (h_t, h'_t, b_t) = \sum_{j,k} \left[ p_{jk} (h_t, h'_t) E[\exp \left( -\varepsilon_{jkt}^*/b_t \right) | h_t] \times \max \left\{ \alpha_{1jk} (h_t)^{1/b_t} E_t \left[ v_{jk,t+1} B (H_{jk} (h_t), H_{jk} (h'_t), b_{t+1}) \right]^{1-1/b_t}, \alpha_{0jk} (h_t)^{1/b_t} E_t \left[ v_{jk,t+1} g (\pi_{j,t+1}, h_t) B \left( H'_{jk} (h_t), H_{jk} (h'_t), b_{t+1} \right) \right]^{1-1/b_t} \right\} \right]
\]

\(B_s (h_t, h'_t, b_t)\) uniformly converges to a unique \(B (h_t, h'_t, b_t)\) as \( s \to \infty \). Defining \(A (h_t, b_t) \equiv B_s (h_t, h'_t, b_t)\) optimal consumption is:

\[
c^*_t = \frac{e_t}{b_t} + \frac{t}{\rho} \ln \beta + \frac{a_t}{\rho b_t} - \rho^{-1} \left\{ \ln E_t \left[ v_{jk,t+1} A (H_{jk} (h_t), b_{t+1}) \right] \right\} + (1 - b_t^{-1}) \varepsilon_{jkt}^* \ln \alpha_{1jk} (h_t)
\]

the optimal job choice indicators \(d_t\) are picked to maximize:

\[
d_{jkt} \varepsilon_{jkt} + \sum_{j=1}^{J} \sum_{k=1}^{K} d_{jkt} \left( \varepsilon_{jkt} - \ln \alpha_{1jk} (h_t) - (b_t - 1) \ln E_t \left[ v_{jk,t+1} A (H_{jk} (h_t), b_{t+1}) \right] \right)
\]

and the equilibrium compensation contract can be expressed as

\[
w_{jk,t+1} (\pi, h_t) = \frac{b_{t+1}}{\rho (b_t - 1)} \ln \alpha_{1jk} (h_t) + \frac{b_{t+1}}{\rho} \ln \left[ A \left( H_{jk} (h_t), b_{t+1} \right) \right] - \frac{b_{t+1}}{\rho} q_{jk} \left[ P (h_t) \right]
\]

\[+ \frac{b_{t+1}}{\rho} \ln \left[ 1 - \eta g_{jk} (\pi, h_t) + \eta \left[ \frac{\alpha_{1jk} (h_t)}{\alpha_{0jk} (h_t)} \right]^{1/(b_t - 1)} \right] \frac{A \left( H_{jk} (h_t), b_{t+1} \right)}{B \left( H'_{jk} (h_t), H_{jk} (h_t), b_{t+1} \right)} \]

where \(P (h_t)\) is defined analogously to above.

5 Identification and Testing

Our model is identified and estimated from longitudinal data on executive compensation, the firm’s abnormal returns, and the transition choices executives make each period conditional on the values of their state variables, factors that affect their current and future payoffs. The primitives of our model are characterized by the probability density function of abnormal returns when every manager is diligent, denoted by \(f_j (\pi)\); the density when only the \(k^{th}\) ranked executive shirks which we express as \(f_j (\pi) g_{jk} (\pi | h_t)\); the coefficient of absolute risk aversion \(\rho\); the parameters affecting tastes for diligent work including the effects of human capital \(\alpha_{1jk} (h)\); the nonpecuniary benefits of shirking \(\alpha_{0jk} (h)\); and the probability distribution of the idiosyncratic disturbance term to preferences \(h (\varepsilon)\). The
Lemma 7

The remaining parameters: $\alpha_{0jk}(h)$, $\alpha_{1jk}(h)$, $h(\varepsilon)$ and $g_{jk}(\pi|h_t)$.

In Section 3 we demonstrated that the minimization problem solved by each firm in our framework has the same structure as the minimization problem a firm would solve if there was no human capital and no alternative job opportunities, that is where $A_s(h) \equiv U_{jk}(h) \equiv 1$. The only difference in the optimal contract distinguishing our model from this specialization emerges from interpreting the mappings $\alpha_{1jk}(h)A_s(h^{jk,1}_{t+1})$ and $U_{jk}(h)$. Gayle and Miller (2008c) analyze identification and estimation of pure moral hazard models where $A_s(h) \equiv U_{jk}(h) \equiv 1$. Given a regularity condition on $g_{jk}(\pi|h_t)$, they show that both $g_{jk}(\pi|h_t)$ and $\alpha_{0jk}(h)$ are exactly identified if and only if $\rho$ and $\alpha_{1jk}(h)$ are identified from the competitive selection equations. Restating their results for our framework, $\alpha_{0jk}(h)$ and $g_{jk}(\pi|h_t)$ are exactly identified if and only if $\alpha_{1jk}(h)$, $h(\varepsilon)$ and $\rho$, and hence $A_s(h)$ and $U_{jk}(h)$, are identified from the competitive selection equations. Thus establishing identification in this framework amounts to proving $\alpha_{1jk}(h)$, $h(\varepsilon)$ and $\rho$ are identified from the competitive selection equations.

In this section we formalize the arguments given in the previous paragraph. Following Gayle and Miller (2008c) we begin by proving that $g_{jk}(\pi|h_t)$ and $\alpha_{0jk}(h)$ are identified if $\rho$ and $\alpha_{1jk}(h)$ are known, the former from the estimated compensation schedule, the latter from the incentive compatibility condition. Then we extend their results by showing that $\alpha_{1jk}(h)$ and $\rho$ are identified providing one of the executive types faces a genuine choice between two different jobs at some point during his life cycle.

5.1 Parameters Characterizing Shirking

To estimate $g_{jk}(\pi|h_t)$ we impose a regularity condition that for all $(j, k, h)$ there exists some finite return $\pi$ such that $g_{jk}(\pi'|h) = 0$ for all $\pi' > \pi$. This assumption implies that, should the firm performance at the end of the period be truly outstanding, then shareholders would be certain that all the executives had worked diligently during the period. Appealing to Proposition 3.1 of Gayle and Miller (2008c), we can then prove $g_{jk}(\pi|h_t)$ is identified if $\rho$ is known. Substituting the expression derive for $g_{jk}(\pi|h_t)$ into Since the incentive compatibility condition yields an expression for $\alpha_{0jk}(h)/\alpha_{1jk}(h)$ in terms of $A_{s-1}(h)$, $v_{k,j,t+1}(\pi, h_t)$ and $g_{jk}(\pi|h_t)$, it immediately follows from their respective definitions and the expression for $g_{jk}(\pi|h_t)$ that if $\alpha_{1jk}(h)$, $h(\varepsilon)$ and $\rho$ are known, then $\alpha_{0jk}(h)$ can be inferred.

Lemma 7

$$g_{jk}(\pi|h_t) = \frac{v_{k,j,t+1}(\pi, h) - v_{k,j,t+1}(\pi, h)}{v_{k,j,t+1}(\pi, h) - E_t[v_{k,j,t+1}(\pi, h)|h]}$$

$$\frac{\alpha_{0jk}(h)}{\alpha_{1jk}(h)} = \left\{ \frac{E_t\{B_{s-1}[H_{jk}(h_t), H_{jk}(h_t'), b_{t+1}]\}}{E_t\{B_{s-1}[H_{jk}(h_t'), H_{jk}(h_t'), b_{t+1}]\}} \right\}^{1-b_t}$$

(10)

The form of the shirking parameter estimates take in (10) suggests a test that differentiates the scenarios we have entertained about whether human capital is observed or not.
If human capital is unobserved, and in addition a variable denoted $h_{t0}$ enters $A_{s-1}(h)$ but not $\alpha_{0jk}(h) / \alpha_{1jk}(h)$, then from (10), variation in $h_{t0}$ does not affect the quantity:

$$\begin{vmatrix}
E_t \{ B_{s-1} [H_{jk}(h_t), H_{jk}(h_t'), b_{t+1}] \} \\
E_t \{ B_{s-1} [H'_{jk}(h_t), H_{jk}(h_t'), b_{t+1}] \}
\end{vmatrix}
$$

Accordingly we can test the joint null hypothesis, that human capital is not observed and $\alpha_{0jk}(h) / \alpha_{1jk}(h)$ does not depend on $h_{t0}$, which implies:

$$\text{cov} \left\{ \begin{vmatrix}
E_t \{ B_{s-1} [H_{jk}(h_t), H_{jk}(h_t'), b_{t+1}] \} \\
E_t \{ B_{s-1} [H'_{jk}(h_t), H_{jk}(h_t'), b_{t+1}] \}
\end{vmatrix}, h_{t0} \right\} = 0$$

### 5.2 Parameters Characterizing Diligence

This leaves $\alpha_{1jk}(h), h(\varepsilon)$ and $\rho$ to analyze from the participation equation and the choice probabilities. We now prove by an induction that $\alpha_{1jk}(h)$ and $\rho$ are identified if at least one type of executive faces a choice between lotteries at some stage in his career . . . more to come

### 6 Data

The data for our empirical study was compiled from three sources. From Standard & Poor’s ExecuComp database we extracted records on the job title and compensation of the eight highest paid executives in the S&P 500, Midcap, and Smallcap firms for the years 1992 through 2006 inclusive. Data on the employer firms were supplemented by the S&P COMPSTAT North America database and monthly stock price data from the Center for Securities Research (CRSP) database. We matched the names, birth dates and gender of 16,300 executives from 1800 firms with information in Who’s Who to augmented their records with biographical data. The resulting data set gives us unprecedented access to detailed firm characteristics, including accounting and financial data, along with their managers’ characteristics, namely the main components of their compensation, including pension, salary, bonus, option and stock grants plus holdings, their socio-demographic characteristics, including age, gender, education, and a description of their career history through the five ranks and firms.

This section summarizes the aggregate features of our data set. We present estimates of the distribution of abnormal returns and show how they vary with executive characteristics. We estimate elasticities of compensation with respect to returns and measures of executive experience. Finally we investigate, empirically, how experience and other background variables affect job transitions and thus define the career paths of executives. In this way we describe the variation in the data that supports the identification and estimation of our model of executive compensation and career choice.

### 6.1 Summary Statistics

Most of the characteristics of the executives and firms in our sample require no explanation, but the construction of several variables merit comment. The sample of firms was
initially partitioned into three industrial sectors by GICS code. Sector 1, called primary, includes firms in energy (GICS:1010), materials (1510), industrials (2010,2020,2030), and utilities (5510). Sector 2, consumer goods, comprises firms from consumer discretionary (2510,2520,2530,2540,2550) and consumer staples (3010,3020,3030). Firms in health care (3510,3520), financial services (4010,4020,4030,4040), information technology and telecommunication services (410, 4520, 4030, 4040, 5010) comprise Sector 3, which we call services. In our sample 37 percent of the firms belong to the primary sector, 28 percent to the consumer goods sector, and the remaining 35 percent to the services sector. Firm size was categorized by total employees and total assets, the median firm in each size category determining whether the other firms are called large or small. The sample mean value of total assets is $18.2 billion (2000 US) with standard deviation $76.2 billion, while the sample mean number of employees is 23,659 with standard deviation 65,702.

Table 1 describes the characteristics of management by sector and firm size. Jobs were assigned to a rank using the hierarchy ordering we developed in our work on gender discrimination. At 27 percent, Rank 2 is the most commonly observed rank, which reflects the diversity of promotion schemes across firms. By way of contrast, the top and bottom ranks each only contribute 6 percent to the sample population. The distribution of ranks across the three sectors is roughly independent but small firms, as measured by either assets of employment, have a greater proportion of their executives congregating in the lower ranks, with 30 percent versus 20 in the bottom two ranks. Four measures of experience were included to capture the potential of on-the-job training. Executive experience is the number of years elapsed since the manager was first recorded as one of the top eight paid executives in the sample. Tenure is years spent working at the employee’s current firm. We also tracked the number of moves the manager made throughout his career in different jobs and ranks, as well as the number of moves since becoming an executive. Promotion is a indicator variable for whether the manager was promoted recently or not.

The mean age of executives is almost 54 years with a standard deviation of about 9. Only 4 percent of the sample are female, ranging between 3 percent in the primary sector and 5 percent in the consumer sector. Roughly speaking, formal education is uniformly distributed evenly between bachelor degree or less, professional certification (in accounting or law for example), MBA, some other Master’s degree, and Ph.D. The distribution is approximately independent of firm size and sector, ranging from 15 percent with an MS/MA in the consumer sector to 27 percent in small firms by employee for professionally certified executives.

Tenure in the firm averages about 14 years, about 40 years less than age, with standard deviation of about 11, two years more. The sectors are ranked the same way with respect to age and tenure; similarly firms with small assets have both the oldest executives and the longest tenure. In these respects average age, firm sector and size are almost sufficient statistics for average tenure, giving the deceptive appearance at this level of aggregation that executives within firms follow a well defined career track. Averaging across the sample, there are two rank and/or firm turnover moves per observation, one of which has occurred since acquiring executive status. About one third of executives have been promoted within the last two years.

The most important differences between the executives across firm size and sector

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8See Gayle, Golan and Miller (2008).
relate to their compensation. Regardless of which measure is used, the mean salary and bonus in small firms is about two thirds the mean in large firms, about half the total compensation, with standard deviations about one third smaller. This suggests that similarly named positions in small firms are not comparable to their analogues in large firms and may help explain differences between internal and external transitions.

Summarizing differences across firm type, the consumer sector has the lowest percent of executives with advance degrees and the highest percent of female executives, while the service sector has the lowest average tenure and the highest promotion rate and highest total compensation. Total compensation is roughly twice as large in large firms (using both measures), promotion and turnover rates are greater, tenure is lower, and there are more executives holding MBA degrees.

Table 2 describes the characteristics of executives by rank. The average age between Rank 1 and 3 declines from 60 to 52, but is more or less constant as rank falls off further. Similarly average tenure is roughly constant in the lower and middle ranks at 14 but rises to 15 and 17 for Ranks 2 and 1 respectively. The average gap between Ranks 1 and 3 in executive experience is 6 years. To summarize, relative to the lower ranks, Ranks 1 and 2 are 8 years older, with only 6 years more executive experience and just 2 years more tenure, late bloomers hired by the firm late in their career. Not that they are likely to move more than those who do not reach the top levels; although 8 years older the they average the same number of past moves, before and after becoming an executive.

Females form a very small fraction of the executive sample, and they are not uniformly distributed by rank. By a factor of two to three, females congregate in the lower executive ranks relative to males; 2 percent of the top two ranks are females, while 6 percent of Ranks 5 and 6 are female. With regard to the education background variables, the two most striking features are that there is higher percent (out of total executives in the rank) of executives with MBA degrees in the top 4 ranks, the percent of executive with another Masters degree or a Ph.D. is greater in the bottom there ranks, and there is a larger percent of executives with professional certification in the bottom 4 ranks.

Average total compensation and the salary components rise from Rank 7, are maximized at Rank 2, at levels that are more than twice as high as the corresponding figures for Rank 7, and decline. The salary component for Rank 1 is only eclipsed by Rank 2, but it is an open question whether the total financial compensation package offered for a Rank 1 position is more or less desirable than the offer for a Rank 5 position. Although the average compensation $2.7 million for Rank 2 exceeds the Rank 5 mean by almost $400,000, the standard deviation for the former is more than twice that of the latter. For example, if all compensation variation observed in the data was resolved before an executive accepted a position, implying the standard deviation simply reflects heterogeneity in fixed pay contracts, then there would be many Rank 5 positions that pay better than many Rank 2 positions. Alternatively if all the variation in compensation was resolved after the executive accepted his job, implying the standard deviation is a measure of the income uncertainty, the executive would prefer Rank 5 to Rank 1 position if he was sufficiently

\footnote{We followed Antle and Smith (1985, 1986), Hall and Liebman (1998), Margiotta and Miller (2000) and Gayle and Miller (2008a, 2008b) by using total compensation to measure executive compensation. Total compensation is the sum of salary and bonus, the value of restricted stocks and options granted, the value of retirement and long term compensation schemes, plus changes in wealth from holding firm options, and changes in wealth from holding firm stock relative to a well diversified market portfolio instead.}
risk averse.

6.2 Abnormal returns

We defined the abnormal returns of the firm as the residual component of returns that cannot be priced by aggregate factors the manager does not control. In an optimal contract compensation to the manager might depend on this residual in order to provide him with appropriate incentives, but it should not depend on changes in stochastic factors that originate outside the firm, which in any event can be neutralized by adjustments within his wealth portfolio through the other stocks and bonds he holds. More specifically, letting \( \vartheta_{jt} \) denote the value of the \( j^{th} \) firm at time \( t \), the gross abnormal return attributable to all the executives’ actions is the residual

\[
\pi_{jt} \equiv \left( \vartheta_{jt} + D_{jt} + \sum_{k=1}^{K} w_{jkt} \right) / \vartheta_{jt-1} - x_t
\]

where \( x_t \) is the return on the market portfolio in period \( t \) and \( D_{jt} \) is the dividend. This study assumes that \( \pi_t \) is a random variable that depends on the managers’ efforts in the previous period but, conditional on the effort vector of the executive branch \( \{l_{jkt}\}_{k=1}^{K} \), is independently and identically distributed across both firms and periods.\(^{10}\)

We now show how abnormal returns depends on the experience and the other characteristics of the executives.

6.3 Compensation

We estimated annual excess returns for firms in equation 11 from the data, and then computed, conditional on the state variables, a nonparametric estimator of total compensation from our imputed values compiled from the data, which we assume is the sum of true compensation and independent measurement error. We used Kernel methods to nonparametrically estimate \( w_{2mk}(\pi, h) \), the compensation schedule for diligent work, for each \((m, k, h)\) as:

\[
w_{2mk}^{(N)}(\pi, h) = \frac{\sum_{s=1, s \neq n}^{N} \sum_{t=1}^{T} w_{st} I \{d_{mkst} = 1, h_{st} = h\} \cdot K\left( \frac{d_{mt} - \pi}{\delta_{sN}} \right)}{\sum_{s=1, s \neq n}^{N} \sum_{t=1}^{T} I \{d_{mkst} = 1, h_{st} = h\} \cdot K\left( \frac{d_{mt} - \pi}{\delta_{sN}} \right)}
\]

Table 3 reports OLS and LAD results from regressing how compensation varies with firms’ and executives’ characteristics. The (conditional) level effects are given in the first two columns of estimates, their interactions with abnormal returns in the second two. Controlling for background demographics and tenure more or less leaves intact the qualitative rank ordering on total compensation we found in Table 3. Total compensation to Ranks 6 and 7 differ by a statistically insignificant amount, and then rises with promotion, spiking at Rank 2, compensation to Rank 1 falling between Ranks 3 and 4. In contrast the unconditional means and standard deviations reported in Table 3, however, the results from the regression analysis separate the effects of excess return, which induces

\(^{10}\)In our sample the mean abnormal return is -0.005 with standard deviation 0.6, and we do not reject the null hypothesis that it is uncorrelated with the stock market.
uncertainty to manager’s total compensation, from the background variables that determine observed heterogeneity. Note that Rank 1 is more affected by excess returns than every rank except 2. Thus Rank 1 has a lower (OLS) or the same (LAD) estimated mean and more dependence on abnormal returns than Rank 3, while Rank 2 has a higher mean but more dependence than Rank 3. Therefore Rank 3 offers a superior total compensation package to Rank 1, and for sufficiently risk averse executives, a more attractive compensation package than the Rank 2. Continuing in this vein, dependence on excess returns is declining in the remaining middle or lower ranks.

All the firm size and sector variables have significant coefficients except the OLS estimator of the level effect distinguishing the consumer from service sector. None of the background variables for executives interact significantly in the OLS regression, but almost all have significant level effects irrespective of estimator. A notable exception are the coefficients relating to gender. The OLS estimator indicates that gender has no effect on compensation level or its dependence on abnormal returns, whereas the LAD estimator implies there is a small positive level effect of $91,731 and significantly reduced dependence on abnormal returns, both factors making an executive positions more attractive to females relative to males.

With respect to education the OLS results show, that after controlling for the other observed differences, Ph.D. and MBA graduates earn more than $300,000 in excess of executives with undergraduate degrees only, who earn $386,793 more than those with professional certification only. Compensation is quadratic in age as is the case in wage regressions for many occupations. Tenure, executive experience and the number of past moves have statistically significant effects on compensation but are small and inconsequential in magnitude. More noteworthy is the large estimated sign-on bonus associated with turnover, $551,859 for LAD and $994,989 for OLS.

Overall our results suggest that after controlling for rank and firm type, there are significant returns from acquiring general human capital in formal education, but little from firm specific capital that is measured in terms of tenure within any one job and/or experience acquired at a variety of jobs. Similarly gender is not a useful predictor of wages given the other executive’s and other characteristics and the nature of the job. To summarize, aside from formal education, job transitions and the abnormal returns of their own firms are the main drivers determining how wealthy executives become.

6.4 State Variables and Conditional Choice Probabilities

We denote the state variables relevant for the \( n^{th} \) manager at the time \( t \) by \( h nt \), one of \( h < \infty \) possible characteristics, the ranks by \( r \in \{1, \ldots, R\} \) and the firm types by \( s \in \{1, \ldots, S\} \). In our model \( h nt, t+1 \), the \( n^{th} \) manager’s state variables in the period \( t+1 \), are fully determined by \( h nt \), the type of firm he transitions to, denoted \( s nt \), and his rank next period, \( r nt \), by a mapping \( h nt, t+1 \equiv f (h nt, r nt, s nt) \), which we define in the next section. Our theory models the transition of \( h nt \) to \( h nt, t+1 \) through the competitive equilibrium choices of \((r nt, s nt)\), a stochastic process that generates the data. The structural estimation of our theoretical framework uses as input reduced form estimates of \( P (r nt, s nt | h nt) \), the probability of \((r nt, s nt)\) conditional on \( h nt \).

We report our estimates for the reduced form of our model. Since \( R \) and \( S \) are finite, and we assume \( H \) is a finite set, it follows that in principle cell estimators could be
used to recover $P(r_{nt}, s_{nt} | h_{nt})$. Although our sample size, 59,066, is very large compared with all previous studies of this market, the comprehensive detail that accompanies each observation also greatly magnifies the total number of cells $RSH$, needed to estimate the model, so this procedure is not feasible. For example only 5 percent of the observations in our sample are female, and none of them have doctorates and head small firms. Many smoothing algorithms are asymptotically equivalent. We used multinomial logits to estimate the reduced form, because of their computational tractability in recovering the structural parameters, because the logit estimates are easy to interpret, and because they illustrate how the variation in our data is used to estimate the underlying structure. For expositional convenience we decomposed $P(r_{nt}, s_{nt} | h_{nt})$ into

$$P(r_{nt}, s_{nt} | h_{nt}) = P(r_{nt} | h_{nt}, s_{nt}) P(s_{nt} | h_{nt})$$

and separately estimated $P(s_{nt} | h_{nt})$, the probability of firm type selected as a function of the state variables, from $P(r_{nt} | h_{nt}, s_{nt})$, the selection of rank conditional on both the state variables and also the firm selected.

Table 4 presents our estimates of $P(s_{nt} | h_{nt})$. The columns refer to the type of firm chosen conditional on moving from the current employer, and the state variables are defined by the rows. The omitted (column) choice is to remain employed with the current firm one more period, and the base line (row) category is a college educated Rank 1 executive employed in a firm of type 1.

MBAs go to 7. MSMAs and Ph.D.’s don’t transit as much, as we saw in the previous table. Controlling for other state variables we now also see that no degree executives also do not move as much as the college educated group. Female behave the same as males. Similarly tenure and male have no significant effects on the probability of an external move. Older execs are more likely to leave and conditional on leaving are less likely to go 3 than the other types.

Perhaps the most striking feature of this table is that when executives move they join firms similar to the ones they left, that is defined in terms of sector and size. Furthermore conditional on moving to a firm of different size, they are more likely to join a firm in the same sector as the one they left. Broadly speaking, the bottom rows, referring to the rank of the executive at the beginning of the period, show that highly ranked executives are less likely to move than the lower ranked ones, evident from the fact that the estimated coefficients increase in each row.

The final column of Table 4 reports on the probability of leaving the sample for at least two years and never returning, a condition we call retirement. The higher the rank the less likely the probability of retirement, indicated by the decreasing sequence of coefficients on rank. Possibly for very different reasons, executives and those without formal qualifications are less likely to exit this sample than groups with other formal education. The indicator variable for gender has a far bigger impact than any of the education variables. Mirroring female labor supply more generally, women in this highly select and lucrative market are more likely to withdraw from it than their male colleagues and competitors. Finally there are significant sector differences.

Finally our estimates of $P(r_{nt} | h_{nt}, s_{nt})$ are presented in Table 5. It shows female executives with a doctorate are more likely to select into the bottom rank. The conditional choice probability estimates shed light on the effects of tenure and age. Here we see that,
controlling for all other state variables, last period employer, and this year’s employer as well, Rank 2 executives are in fact older than Rank 1 executives, signified by the higher coefficient estimate. Given values of the other observed factors, lower ranked employees have more tenure. The highest coefficients invariably show staying in the same rank is the most likely outcome, and an executive in the lowest rank is more likely to move to Rank $i$ than Rank $i + 1$. Similarly Rank 4 executives are more likely to be demoted than be promoted to Rank 3, evident from the estimated coefficients in Table 4. The results in Table 4 show that relative to other executives, turnover for a Rank 2 manager is more likely than external promotion.

7 Estimation

The bulk of this section lays out a sequential estimator for the model and reports our estimates of its several components. Given the minimal movement in bond prices over this period, we assumed the bond price is constant, setting $b_t = b$ for all $t$. For computational ease and to reduce the computational burden we estimated the stationary infinite horizon model. The taste and human capital parameters $\rho$ and $\alpha_{1jk}(h)$ are estimated from the participation constraint, exploiting the idea that when risk averse managers make rational choices between different uncertain outcomes or lotteries they are revealing their attitude towards risk. The likelihood ratio $g_{jk}(\pi | h)$ is estimated from the curvature in the optimal compensation schedule. The shirking parameters $\alpha_{0jk}(h)$ are estimated from the incentive compatibility constraint, which reflects the fact that shareholder compensate managers just enough to deter them from engaging in activities that do not maximize the value of the firm.

7.1 Utility and Human Capital

In our estimation we assumed the disturbance is standard Type 1 Extreme Value, and in our initial estimates. Substituting the estimators for $P_{jk}(h)$ and $w_{2jk}^o(h)$ obtained from the previous section into

$$E_t [U_{jk}(h) - \alpha_{1jk}(h) B_{s-1}(h) v_{j,k,t+1}(\pi, h) | h] = 0$$

Our estimator of the $\alpha_{2jk}$ parameters and the $\rho$, based on the competitive selection equations, is $\sqrt{NT}$ consistent and asymptotically normal, the covariance differing from the standard formula only because the choice probabilities and the compensation schedule are estimated in the first two steps. But rather than form JKZ orthogonality conditions from the conditional expectation functions, we formed a GMM estimator from the implied covariances

$$E \{[U_{jk}(\pi) - \alpha_{2jklt} B_{s-1}(h) v_{j,k,t+1}(\pi, h)] | h\} = 0$$

using the counting variables, tenure, executive experience and age as instruments, after substituting in an approximating function $\hat{U}_{jk}(h)$ for $U_{jk}(h)$. The former differs from the latter only because consistent estimators for $P_{jk}(h)$ and $w_{jk}^*(h)$ are used instead of their true values. The remaining background variables, categorical variables signifying educational background and gender, were also used as conditioning variables in forming the orthogonality functions for the estimator. Second, when forming the recursion that
defines \( A_s(h) \), used in the definitions of \( U_{jk}(h) \) and \( \tilde{U}_{jk}(h) \), we exploited the fact that, given the manager’s choice, the transition of \( h_{nt} \) to \( h_{nt+1} \) is deterministic. Using the definition of conditional probability and sufficiency:

\[
\Pr \left( h' | h_{nt}, d_{ijk}, d_{mnt} \right) = \Pr \left[ d_{j',k'_{t+1}} = 1, h_{nt+1} = h' | h_{nt}, d_{jk} = 1, l_{ijk} = 1 \right] \\
= \Pr \left[ d_{j',k'_{t+1}} = 1 | h_{nt+1} = h', h_{nt}, d_{jk} = 1, l_{ijk} = 1 \right] \\
\times \Pr \left[ h_{nt+1} = h' | h_{nt}, d_{j',k'_{t+1}} = 1 \right] \\
= \Pr \left[ d_{j',k'_{t+1}} = 1 | h_{nt+1} = h', h_{nt}, d_{jk} = 1, l_{ijk} = 1 \right] \\
\times I \{ h_{nt+1} = h' | h_{nt}, d_{j',k'_{t+1}} = 1 \}
\]

### 7.2 The Remaining Parameters

We used nonparametric methods to recover

\[
g_{jk}(\pi|h) = \frac{v^{-1}_{k,j,t+1}(\pi,h) - v^{-1}_{k,j,t+1}(\pi,h)}{v^{-1}_{k,j,t+1}(\pi,h) - E_t[v_{k,j,t+1}(\pi,h)|h]}
\]

and:

\[
\alpha_{0jk}(h) = \alpha_{1jk}(h) E_t \left\{ B_{s-1} \left[ H_{jk}(h_t), H_{jk}(h'_t), b_{t+1} \right] \right\} \left[ g_{jk}(\pi|h) \right]^{1-b_t}
\]

### 8 Investment versus Moral Hazard

In the concluding section to this paper we assess how much agency problems in executive markets are mitigated by their career concerns. Two of the four metrics we use measure the impact of an executive shirking rather than working. We estimated how much abnormal returns would fall if shareholders failed to incentivize one of its executives but continued to pay the other according to the optimal schedule. This is one measure of how much a firm stands to lose by ignoring the moral hazard problem. The executive, on the other hand, is much more concerned with the compensating differential between diligence and shirking. We computed the compensating differential to an executive from following his interests (shirking) rather than acting according to the interests of the shareholders (working diligently). The other two metrics focus on the cost of eliminating the moral hazard problem. We report on how much the firm pays to induce diligence in the presence of human capital investment, a risk premium for eliminating the moral hazard problem. Finally we calculate how much more a firm would have to pay if executives were not motivated by career concerns, ambition that helps to internalize what would otherwise be a more substantial moral hazard problem.

Each metric was computed using the structural estimates obtained from the previous section, by executive rank, averaged over firm type and executive background. Thus successive rows in Table 6 report a sample average for the rank and its standard deviation, conditional on optimal behavior by the rest of the management team. For the purposes of comparisons with other studies in this literature we also report the estimated risk aversion parameter, the top entry. Quite plausible, and comparable to previous estimates found, we note that an executive with exponential utility and risk aversion parameter of 0.45
would be willing to pay $217, 790 to insure against an actuarially fair gamble that offers a loss of $1 million with probability one half and a gain of $1 million with probability one half.

The first metric is an average over $\tau_{1mk}(h)$, the expected gross loss in the value of the firm of type $m$ in percentage terms if a rank $k$ executive with background $h$ tends his own interests for one year, instead of maximizing the expected value of the firm, that is before netting out the decline in expected compensation all executives would incur from the deteriorating financial performance of the firm. When all executives work diligently, by definition abnormal returns have mean zero, meaning $E[\pi] = 0$. Thus $\tau_{1mk}(h)$ is found by integrating abnormal returns conditional on the executive in question shirking, when every other executive works diligently:

$$\tau_{1mk}(h) \equiv E \{ \pi [1 - g_{mk}(\pi, h)] \} = -E[\pi g_{mk}(\pi, h)]$$

We interpret $\tau_{1mk}(h)$ as a measure of the executive’s span of control, because it indicates his potential impact on the firm from behaving irresponsibly. Not surprisingly we find Rank 2 executives exercise the greatest span of control; at 11 percent per year, a chief executives can drive the value of firm equity down to less than half its current value in 8 years, shareholders willing. Similarly, the result that the estimated span of control declines through the middle and lower ranks, confirms our intuition. More remarkable is our finding that executives in Ranks 2 and 3 have a greater span of control than those in Rank 1, as do many in Rank 4.

Taking the manager’s perspective rather than the firm’s, the compensating differential between working hard and shirking, which we denote by $\tau_{2mk}(h)$, is measured by differencing $w^0_{1mk}(h)$, the manager’s reservation certainty equivalent wage to shirk, from $w^0_{2mk}(h)$, the manager’s reservation certainty equivalent wage to work diligently under perfect monitoring. Derived from the participation constraint, these certainty equivalents can be expressed as:

$$w^0_{1mk}(h) = \frac{b_t+1}{\rho} \log(\alpha^{t+1,1}_{mkt}(h)) + \frac{b_t+1}{\rho(b_t-1)} \log(\alpha_{1mk}/U^E_{mk}(h_m))$$

and

$$w^0_{2mk}(h) = \frac{b_t+1}{\rho} \log(\alpha^{t+1,2}_{mkt}(h)) + \frac{b_t+1}{\rho(b_t-1)} \log(\alpha_{2mk}/U^E_{mk}(h_m))$$

Thus

$$\tau_{2mk}(h) \equiv w^0_{2mk}(h) - w^0_{1mk}(h) = \frac{b_t+1}{\rho} \log(\alpha^{t+1,2}_{mkt}(h)/\alpha^{t+1,1}_{mkt}(h)) + \frac{b_t+1}{\rho(b_t-1)} \log(\alpha_{2mk}/\alpha_{1mk})$$

If a manager does not maximize the value of the firm, he gains utility from the nonpecuniary benefits of pursuing his own interests, but does not acquire so much human capital, and thus reduces his chances of higher wages and better positions in the future.

The first factor would also arise in a static model of pure moral hazard where there are no career concerns, and in our formulation does not depend on the executives background characteristics:

$$\tau^PM_{2mk} \equiv \frac{b_t+1}{\rho(b_t-1)} \log(\alpha_{2mk}/\alpha_{1mk})$$
Our estimates in Table 6 show that contemporaneous nonpecuniary shirking/working benefit differential associated with the Rank 2 position, at $2.48 million, exceed those associated with any of the other ranks, but that the annual differential from the Rank 1 position is the next highest. Thus Rank 1 has a lesser span of control than Rank 3, but more nonpecuniary benefits. Again these benefits decline through the middle and lower ranks.

The second factor determining $\tau_{2mk}(h)$ reflects those dynamic features of our framework relating to career concerns

$$\tau^H_{2mk}(h) \equiv \frac{b_{t+1}}{\rho} \log(\alpha_{mkt}^{t+1.2}(h)/\alpha_{mkt}^{t+1.1}(h))$$

Here we find that, on average, the benefits of human capital accumulation decline monotonically with rank, and that compared with $\tau^{PM}_{2mk}$, are much less dispersed throughout the population of firm types and executive backgrounds. At the lower ranks these benefits are quite considerable. On average a Rank 5 executive is willing to forego $1.88 million per year because of the greater opportunities working diligently versus shirking affords him, while a Rank 1 executive only values the human capital component of the compensating differential at $400,000 million per year.

By inspection the compensating differential $\tau_{2mk}(h)$ is the sum of these two factors

$$\tau_{2mk}(h) = \tau^H_{2mk}(h) + \tau^{PM}_{2mk}$$

Our estimates imply the compensating differential for every rank except the second is about $2 million per year, but exceeds $3 million per year for Rank 2 executives.

How much a firm would be willing to eliminate moral hazard is measured by $\tau_{3mk}(h)$. Under a perfect monitoring scheme shareholders would pay a manager the fixed wage of $w_{2mk}(h)$, and thus eliminate the risk premium they pay him in the form of a favorable lottery over the outcome of abnormal returns to induce diligent work. Hence the expected value of a perfect monitor to shareholders, denoted $\tau_{3mk}(h)$, is the difference between expected compensation under the current optimal scheme and $w_{2mk}^0(h)$, or:

$$\tau_3 \equiv E[w_{mk}(\pi) | h] - w_{2mk}^0(h)$$

$$= E[w_{mk}(\pi) | h] - \frac{b_{t+1}}{\rho} \log(\alpha_{mkt}^{t+1.2}(h)) - \frac{b_{t+1}}{\rho(b_t - 1)} \log(\alpha_{2mk}/U_{mk}^E(h_m))$$

Our findings in Table 6 show that the firms are prepared to pay hardly anything to eliminate the moral hazard problem at the lower ranks, but that at the Ranks 1 and 3, the benefits of a perfect monitor are considerably more. Curiously, the average risk premium paid to Ranks 1 and 3, $1.6$ million and $1.7$ million respectively, are quite close, despite the fact that the other measures of moral hazard are not.

As one final check on the relevance of human capital to resolving moral hazard problems in the executive market, we estimated the extra premium shareholders would pay to eliminate the moral hazard problem if the benefits of acquiring human capital was ignored by an executive, say because neither the organizational structure nor the market rewarded his diligence. In our model this is represented by:

$$\tau_{4mk}(h) \equiv \frac{b_{t+1}}{\rho} \log(\alpha_{mkt}^{t+1.2}(h))$$
The estimates in Table 6 show that career concerns greatly ameliorate the moral hazard problem for lower level executives but their importance declines monotonically with promotion through the ranks, bordering on irrelevance for many Rank 1 executives.

9 Appendix

Proof of Lemma 1. For the metric space $B(h)$ we define a distance metric $d(B_0(h), B_1(h))$ and its norm $\|B_0(h)\|$, and without loss of generality rescale utility by choosing the taste parameters $\alpha_{jkl}$ such that:

$$\left\|\alpha_{jkl}^{1/b_t} E[\exp(-\epsilon_{jkt}/b_t) | h_t] \{E[v_{jkt+1}|h_t, l_t]\}^{1-\frac{1}{b_t}}\right\| \leq k$$

for some real number $k \in (0,1)$. Defining the operator:

$$\Gamma[B(h_t)] \equiv \sum_{(j,k,l)} p_{jkl}(h_t) \alpha_{jkl}^{1/b_t} \left[ B_0 \left( h_{t+1}^{(j,k,l)} \right) \right]^{1-\frac{1}{b_t}} E[\exp(-\epsilon_{jkt}/b_t) | h_t] \{E[v_{jkt+1}|h_t, l_t]\}^{1-\frac{1}{b_t}}$$

it follows that:

$$d(\Gamma[B_0(h_t)], \Gamma[B_1(h_t)])$$

$$= \left\| \sum_{(j,k,l)} p_{jkl}(h_t) \alpha_{jkl}^{1/b_t} \left[ B_0 \left( h_{t+1}^{(j,k,l)} \right) \right]^{1-\frac{1}{b_t}} E[\exp(-\epsilon_{jkt}/b_t) | h_t] \{E[v_{jkt+1}|h_t, l_t]\}^{1-\frac{1}{b_t}} \right\|$$

$$= \left\| \sum_{(j,k,l)} p_{jkl}(h_t) \alpha_{jkl}^{1/b_t} \left[ B_0 \left( h_{t+1}^{(j,k,l)} \right) \right]^{1-\frac{1}{b_t}} E[\exp(-\epsilon_{jkt}/b_t) | h_t] \{E[v_{jkt+1}|h_t, l_t]\}^{1-\frac{1}{b_t}} \right\|$$

$$\leq \sum_{(j,k,l)} \alpha_{jkl}^{1/b_t} \left[ B_0 \left( h_{t+1}^{(j,k,l)} \right) \right]^{1-\frac{1}{b_t}} \left[ \left( B_0 \left( h_{t+1}^{(j,k,l)} \right) \right)^{1-\frac{1}{b_t}} - \left[ B_1 \left( h_{t+1}^{(j,k,l)} \right) \right]^{1-\frac{1}{b_t}} \right]$$

$$\leq k \left\| \sum_{(j,k,l)} \left[ B_0 \left( h_{t+1}^{(j,k,l)} \right) \right]^{1-\frac{1}{b_t}} - \left[ B_1 \left( h_{t+1}^{(j,k,l)} \right) \right]^{1-\frac{1}{b_t}} \right\|$$

$$\leq kd \left( \left[ B_0 \left( h_{t+1}^{(j,k,l)} \right) \right]^{1-\frac{1}{b_t}}, \left[ B_1 \left( h_{t+1}^{(j,k,l)} \right) \right]^{1-\frac{1}{b_t}} \right)$$

$$\leq kd \left( \left[ B_0 \left( h_{t+1}^{(j,k,l)} \right) \right], \left[ B_1 \left( h_{t+1}^{(j,k,l)} \right) \right] \right)$$

This proves $\Gamma$ is a contraction operator. By Banach’s theorem there exists a unique fixed point which we denote by $A(h_t)$, the limit of $\Gamma^*[B_0(h_t)]$ for any starting point $B_0(h_t)$.

Setting $B_0(h_t) = 1$ and $A_s(h_t) = \Gamma^*[B_0(h_t)]$ proves the lemma. ■
Proof of Lemma 2. For all \( s \in \{1, \ldots, T\} \) we set \( A_0 (h_t) \equiv 1 \) for all \((j,k)\) and recursively define \( A_s (h_t) \) as:

\[
A_s (h_t) = -b_t \sum_{(j,k,l)} P_{jk} (h_t) \alpha_{jkt}^{1/b_t} \left[ A_{s-1} (h_{t+1}) \right]^{1 - \frac{1}{n_t}} E \{ \exp (-\varepsilon^*_{jkt} / b_t) | h_t \} \{ E[v_{jk,t+1} | h_t, l_t] \}^{1 - \frac{1}{n_t}}
\]

where \( \varepsilon^*_{jkt} \) is the value of the period \( t \) disturbance when \((j,k)\) is selected, \( b_t \) is the current price of a perpetual bond at \( t \), and \( h_{t+1}^{(j,k,l)} \) is the value of the state variables in period \( t + 1 \) induced by the choices \((j,k,l)\) in period \( t \). From Proposition 1 of Margiotta and Miller (2000, page 678), the value function solving the consumption savings problem at retirement date \( T + 1 \) is:

\[
V_0 (h) \equiv -b_{T+1} \exp \left[ - (a_{T+1} + \rho e_{T+1}) / b_{T+1} \right]
\]

where \((a_{T+1}, e_{T+1})\) are defined in the text. Suppose a manager works in firm and rank coordinate pair \((m,k)\) at time \( T \) for one period and then retires. After selecting \((m,k)\) he chooses consumption and next period’s endowment \((e_T, e_{T+1})\) optimally to maximize:

\[
-\alpha_{jkt} \exp (-\varepsilon^*_{jkt}) \exp (-\rho e_T) - E_T \left[ b_{T+1} \exp \left( -\frac{a_{T+1} + \rho e_{T+1}}{b_{T+1}} \right) v_{jk,T+1} | h_T, l \right]
\]

subject of his budget constraint. From Equation (15) of Margiotta and Miller (2000, page 680), the solution to this problem yields a value function of:

\[
V_{jklT} (h_T) \equiv -b_T \alpha_{jklT}^{1/b_T} \left[ A_0 (h_{T+1}^{(j,k,l)}) \right]^{1 - \frac{1}{n_T}} \exp \left( -\varepsilon^*_{jkt} / b_T \right) \{ E[v_{jk,T+1} | h_T, l] \}^{1 - \frac{1}{n_T}} \exp \left( -\frac{a_T + \rho e_T}{b_T} \right)
\]

Integrating over \((\varepsilon_{11T}, \ldots, \varepsilon_{MKT})\), the idiosyncratic disturbance vector that is revealed at the beginning of the period, and the resulting job and effort level choice \((j,k,l)\) yields

\[
V_T (h_T) \equiv -b_T \sum_{(j,k,l)} P_{jk} (h_T) V_{jklT} (h_T)
\]

\[
= -\exp \left( -\frac{a_T + \rho e_T}{b_T} \right) b_T \sum_{(j,k,l)} P_{jk} (h_T) \alpha_{jklT}^{1/b_T} \left[ A_0 (h_{T+1}^{(j,k,l)}) \right]^{1 - \frac{1}{n_T}} E \left[ \exp (-\varepsilon^*_{jkt} / b_T) | h_T \right] \{ E[v_{jk,T+1} | h_T, l] \}^{1 - \frac{1}{n_T}}
\]

\[
= -\exp \left( -\frac{a_T + \rho e_T}{b_T} \right) A_1 (h_T)
\]

The proof is completed with an induction showing that for all \( s \in \{1, \ldots, T-1\} :\)

\[
V_{jkl} (h_t) \equiv -b_t \alpha_{jkl}^{1/b_t} \left[ A_{s-1} (h_{t+1}^{(j,k,l)}) \right]^{1 - \frac{1}{n_t}} \exp \left( -\varepsilon^*_{jkt} / b_t \right) \{ E[v_{jk,t+1} | h_t, l_t] \}^{1 - \frac{1}{n_t}} \exp \left( -\frac{a_t + \rho e_t}{b_t} \right)
\]

and

\[
V_{T-s+1} (h_t) = -\exp \left( -\frac{a_t + \rho e_t}{b_t} \right) A_s (h_t)
\]
Suppose the equation is true for all $r \in \{1, \ldots, s\}$ for $s < T - 1$. Then the solution to the consumption savings decision at time period $t = T - s - 1$ is found by maximizing:

$$-\alpha_{jklt} \exp (-\varepsilon^*_jkt) \exp (-\rho c_t) - A_{s-1}\left(h_{t+1}^{(j,k,l)}\right) E_t[b_{t+1} \exp \left(-\frac{a_{t+1} + \rho e_{t+1}}{b_{t+1}}\right) v_{jkt, t+1} | h_t, l_t]$$

$$= -\alpha_{jklt} \exp (-\varepsilon^*_jkt) \exp (-\rho c_t) - E_t[b_{t+1} \exp \left(-\frac{a_{t+1} + \rho e_{t+1}}{b_{t+1}}\right) A_{s-1}\left(h_{t+1}^{(j,k,l)}\right) v_{jkt, t+1} | h_t, l_t]$$

with respect to $(c_t, e_{t+1})$. Substituting $t$ for $T$ and $A_{s-1}\left(h_{t+1}^{(j,k,l)}\right) v_{jkt, t+1}$ for $v_{jkt, T+1}$ in Equation above follows directly. Integrating over $(\varepsilon_{1lt}, \ldots, \varepsilon_{JKl})$, the idiosyncratic disturbance vector that is revealed at the beginning of the period, and the resulting job and effort level choice $(j, k, l)$ yields:

$$V_{T-s+1}(h_t) \equiv \sum_{(j,k,l)} P_{jkt, T-s+1}(h_t) V_{jkl}(h_t)$$

$$= -\exp \left(-\frac{a_t + \rho e_t}{b_t}\right) b_t \sum_{(j,k,l)} P_{jkt, T-s+1}(h_t) \alpha_{jklt}^{1/b_t} A_{s-1}\left(h_{t+1}^{(j,k,l)}\right)$$

$$E \left[\exp (-\varepsilon^*_jkt/b_t) | h_t\right] \{E[v_{jkt, t+1} | h_t, l_t]\}^{1-\frac{1}{b_t}}$$

$$= -\exp \left(-\frac{a_t + \rho e_t}{b_t}\right) A_s(h_t)$$

as required, the third line following from the recursive definition of $A_s(h_t)$. \[\blacksquare\]

**Proof of Lemma 3.** For notational convenience define:

$$W_{jkt} \equiv \log \alpha_{jklt} + (b_t - 1) \log \left[A_{s-1}\left(h_{t+1}^{(j,k,l)}\right)\right] + (b_t - 1) \log \left\{E[v_{jkt, t+1} | h_t, l_t]\right\}$$

Then $(j, k)$ is chosen if:

$$\varepsilon_{jkt} + W_{jkt} \geq \varepsilon_{j'k'l} + W_{j'k'l}$$

for $l = l_t$. Let $G(\varepsilon_{1lt}, \ldots, \varepsilon_{JKl})$ denote the probability distribution function for $(\varepsilon_{1lt}, \ldots, \varepsilon_{JKl})$ and $G_{jk}(\varepsilon_{1lt}, \ldots, \varepsilon_{JKl})$ its derivative with respect to $\varepsilon_{jkt}$. Since $G(\varepsilon_{1lt}, \ldots, \varepsilon_{JKl})$ is the product of independently distributed standard Type 1 Extreme value probability distributions in our model:

$$G_{jk}(\varepsilon_{1lt}, \ldots, \varepsilon_{JKl}) = \exp (-\varepsilon_{jkt}) \prod_{(j', k')} \exp \left[-\exp (-\varepsilon_{j'k'l})\right]$$

Using the well known fact that:

$$W_{jkt} - W_{j'k'l} = \log p_{jkt} - \log p_{j'k'l}$$

it now follows that:

$$G_{jk}(\varepsilon_{1lt} + W_{jkt} - W_{11lt}, \ldots, \varepsilon_{jkt} + W_{jkt} + W_{JKl})$$

$$= \exp (-\varepsilon_{jkt}) \prod_{(j', k')} \exp \left[-\exp (-\varepsilon_{jkt} + W_{j'k'l} - W_{jkt})\right]$$

$$= \exp \left[ -\varepsilon_{jkt} - \sum_{(j', k')} \exp (-\varepsilon_{jkt} + W_{j'k'l} - W_{jkt}) \right]$$

$$= \exp \left[ -\varepsilon_{jkt} - \exp (-\varepsilon_{jkt}) \left\{ \sum_{(j', k')} \exp (\log p_{j'k'l} - \log p_{jkt}) \right\} \right]$$

$$= \exp \left[ -\varepsilon_{jkt} - \exp (-\varepsilon_{jkt} - \log p_{jkt}) \left\{ \sum_{(j', k')} \exp (\log p_{j'k'l}) \right\} \right]$$

$$= \exp \left[ -\varepsilon_{jkt} - \exp (-\varepsilon_{jkt} - \log p_{jkt}) \right]$$
From Equation the conditional choice probability for \((j, k)\) can be expressed as

\[ p_{jkt} = \int_{-\infty}^{\infty} G_{jk} (\varepsilon_{jkt} + W_{jktl} - W_{11lt}, \ldots, \varepsilon_{jkt} + W_{jktl} + W_{JKlt}) \, d\varepsilon_{jkt} \]

Hence the probability density function of \(\varepsilon^*_{jkt} \equiv d_{jk} \varepsilon_{jkt}\) is Type 1 extreme value with location parameter \(-\log p_{jkt}\) and unit scale parameter since:

\[
h (\varepsilon^*_{jkt}) = p_{jkt}^{-1} \frac{\partial}{\partial \varepsilon^*_{jkt}} \left[ \int_{-\infty}^{\varepsilon^*_{jkt}} G_{jk} (\varepsilon_{jkt} + W_{jktl} - W_{11lt}, \ldots, \varepsilon_{jkt} + W_{jktl} + W_{JKlt}) \, d\varepsilon_{jkt} \right]
\]

\[
= p_{jkt}^{-1} \exp \left[ -\varepsilon^*_{jkt} - \exp \left( -\varepsilon^*_{jkt} - \log p_{jkt} \right) \right]
\]

\[
= \exp \left[ -\varepsilon^*_{jkt} - \log p_{jkt} - \exp \left( -\varepsilon^*_{jkt} - \log p_{jkt} \right) \right]
\]

To derive:

\[
E \left[ \exp \left( \frac{\varepsilon^*_{jkt}}{b_t} \right) \right]
\]

we draw from Equations (15) and (17) of Chapter 21 of Johnston and Kotz (1970, pages 277 - 278) proving that the moment generating function for \(\varepsilon^*_{jkt}\) is:

\[
E \left[ \exp \left( t \frac{\varepsilon^*_{jkt}}{b_t} \right) \right] = \exp \left( -t \log p_{jkt} \right) \Gamma \left( 1 - t \right)
\]

Setting \(t = \frac{b_t^{-1}}{b_t}\) this simplifies to:

\[
E \left[ \exp \left( \frac{\varepsilon^*_{jkt}}{b_t} \right) \right] = \exp \left( -\log p_{jkt}^{1/b_t} \right) \Gamma \left( \frac{(b_t - 1)}{b_t} \right) = p_{jkt}^{-1/b_t} \Gamma \left( \frac{(b_t - 1)}{b_t} \right)
\]

Proof of Lemma 5. Let \(\Psi\) denote the choice mechanism of the executive, a mapping from wages into choice probabilities derived from the discrete problem the executive solves. Thus \(p = \Psi [w(h)]\). Let \(\Omega\) denote the wage mechanism of the firm, derived from the cost minimization problem it solves. Thus \(w(h) = \Omega (P)\). Form the composite operator \(\Gamma [w(h)] \equiv \Omega \{\Psi [w(h)]\}\). We seek to show \(\Gamma [w(h)]\) has a fixed point, or that there exists a wage function \(w^* (h)\) satisfying \(w^* (h) \equiv \Gamma [w^* (h)]\). To prove \(\Gamma [w(h)]\) has a fixed point we apply a standard fixed point theorem. Hence from the definition of \(\Gamma\) and \(w^* (h)\) it now follows that \(w^* (h) \equiv \Omega \{\Psi [w^* (h)]\}\). Thus:

\[
P \equiv \Psi [w^* (h)] \equiv \Psi (\Omega \{\Psi [w^* (h)]\}) \equiv \Psi [\Omega (P)]
\]

as required by the lemma. ■
References


Table 1: Executives Characteristics by Sector and Firm Size
Compensation and Salary are measured in Thousand of 2006US$

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<tr>
<th>Variable</th>
<th>Service</th>
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<th>Consumer</th>
<th>Asset Small</th>
<th>Asset Large</th>
<th>Employee Small</th>
<th>Employee Large</th>
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*Standard Deviation in Parenthesis*
Table 2: Executives Characteristics
Compensation and Salary are measured in Thousand of 2006 US$

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Table 3: Compensation Regressions

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Table 3 (cont.): Compensation Regressions

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<td>(312.742)</td>
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<td>(312.742)</td>
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<td>(854.534)</td>
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Table 4: Multinominal Logit of Firm Choice  
(Staying with your Current Firm in the Based)

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<tr>
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<td>0.340</td>
<td>0.165</td>
<td>0.061</td>
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Table 6: Structural Estimates and Simulations

$\tau_2$, $\tau_3$, and $\tau_4$ are measured in US$100,000$ of dollars

$\tau_1$ is measured in percentage per year

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